

DYNAMIC SCREENING

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ABSTRACT:

We study dynamic screening problems where elements are subjected to noisy evaluations and, in every stage, some of the elements are rejected while the remaining ones are independently re-evaluated in subsequent stages. We prove that, *ceteris paribus*, the quality of a dynamic screening process is not monotonic in the number of stages. Specifically, we examine the accepted elements' values and show that adding a single stage to a screening process may produce inferior results, in terms of stochastic dominance, whereas increasing the number of stages substantially leads to a first-best outcome.

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1 Introduction

For this paper to get published, it must go through a dynamic screening process. First, the paper is inspected by the administrative staff to verify that it fits the journal's requirements. Then, assuming the paper does fit the technical requirements, the editor and co-editor decide whether to send it out for review. Lastly, it (hopefully) reaches the reviewers for consideration and comments. As the paper is prone to disqualification in any given stage, this process is essentially a dynamic screening. Such dynamic processes are also used in employee-requirement procedures, whether it is in the private sector, government, or academia, and they are also vital in many decision problems where various possibilities are first reduced to a smaller sample, before the final choice is made. One can even find dynamic screening in data analysis (namely, data cleansing) and numerical computations.

In this paper, we study the problem of dynamic screening by focusing on a decision maker who screens elements from a general set, based on noisy valuations. The screening process could vary from a single stage to multiple ones, while keeping an overall capacity constraint on accepted elements. In other words, once a dynamic screening process is in place, a subset of the elements are rejected in every stage, while the remaining ones are independently re-evaluated in subsequent stages.

In this framework, we strive to identify the key advantages and disadvantages of dynamic screening. We follow this goal through two complementing paths. The first is a comparison of a one-stage and a two-stage screening. The second path follows an asymptotic approach, dealing with multi-stage screening where the number of stages increases. The joint results from these two research paths shed light on the questions of why do people use dynamic screening, and when do additional screening stages lead to suboptimal outcomes.

Specifically, the first research path introduces a basic concept that we refer to as *the hidden cost of dynamic screening*. At face value, the introduction of additional screening stages has a direct cost for the time and the effort of performing more tests. The basic intuition is that the superior expected outcomes over-compensate for the added cost. However, we show that additional stages may, in fact, generate *worse* results, both in terms of expected value and in terms of stochastic dominance, even if the accuracy of the additional stages is superior. The basic idea behind this result is that, given a capacity constraint on accepted elements, the introduction of additional stages must be accompanied by some *self-induced slack*. One cannot simply add screening stages without either lowering the bar in preliminary stages and introducing suboptimal elements, or otherwise violating the capacity constraint. Once these suboptimal elements are introduced, the noise in subsequent stages becomes more effective/destructive, although it is strictly more informative (in the sense of Lehmann (1988), for example) compared to previous stages. Let us explain this insight through the following (stylized)

example, which will also serve to present the key findings from our second research method.¹

1.1 The hidden cost of dynamic screening: an example of self-induced slack

Assume that the decision maker is a team leader (TL) who interviews applicants in some big tech company. To simplify the exposition, assume that the suitability of applicants for the available position is given by a random variable $V \sim U[0, 1]$. Since interviews and CVs are rather “noisy”, the TL cannot observe the realized value v of every applicant, but only $v + N_1$, where $N_1 \sim U[-1/4, 1/4]$ is an additive unbiased noise variable (i.e., independent and symmetric about zero). Moreover, the number of open positions dictates an overall acceptance rate of 5%. Following Lagziel and Lehrer (2020), it is optimal for the TL to use a threshold strategy. Thus, to meet the capacity constraint, she fixes a threshold so that a realized noisy valuation is accepted if and only if it is above that threshold. This one-stage screening (process) yields a conditional distribution over the accepted applicants’ suitability, denoted V_1 , whose CDF $F_{V_1}(v)$ is given by the solid (blue) graph in Figure 1(a).

One day, the company decides to improve its screening process by introducing an additional test: an interview with the CTO. The CTO has better screening abilities, characterized by a noise variable $N_2 \sim U[-1/5, 1/5]$. Due to the CTO’s limited time and given the capacity constraint, the TL’s initial screening provides a short list of the top 10% of applicants (according to the TL’s evaluation). Then the CTO performs the additional screening to meet the 5% constraint. This two-stage screening yields a conditional distribution of values, denoted V_2 , whose CDF $F_{V_2}(v)$ is given by the dashed (red) graph in Figure 1(a).

Computing the conditional expected value for each process, we get that $\mathbb{E}[V_1] > \mathbb{E}[V_2]$. So not only the extra screening step is costly, it actually generates an inferior outcome. Moreover, in case the overall acceptance rate is 3% and assuming the CTO maintains the 2 : 1 acceptance ratio, i.e., the TL provides a short list of the top 6% based on the TL’s noisy evaluations, then the one-stage screening *first-order stochastically dominates* the two-stage process [see the respective graphs in Figure 1(b)].

The reason for this phenomenon is the *self-induced slack* of dynamic screening. The introduction of additional screening stages must be accompanied by a reduced acceptance criterion (i.e., lower threshold) in preliminary stages. This transfers some additional burden into advanced stages, where the noise and the sample space reduce *simultaneously*, which can largely affect the outcome. In other words, there is a built-in trade-off where the sample size and noise decrease at the same time, so the

¹The diligent reader may have already realized that our framework and results extend (and, in some cases, are equivalent) to a single-stage multi-evaluators screening, where multiple independent referees conduct an evaluation at the same time, as in the academic peer-review process. The example in Section 1.1 relates to such equivalent settings, as well.

One-stage versus two-stage screening

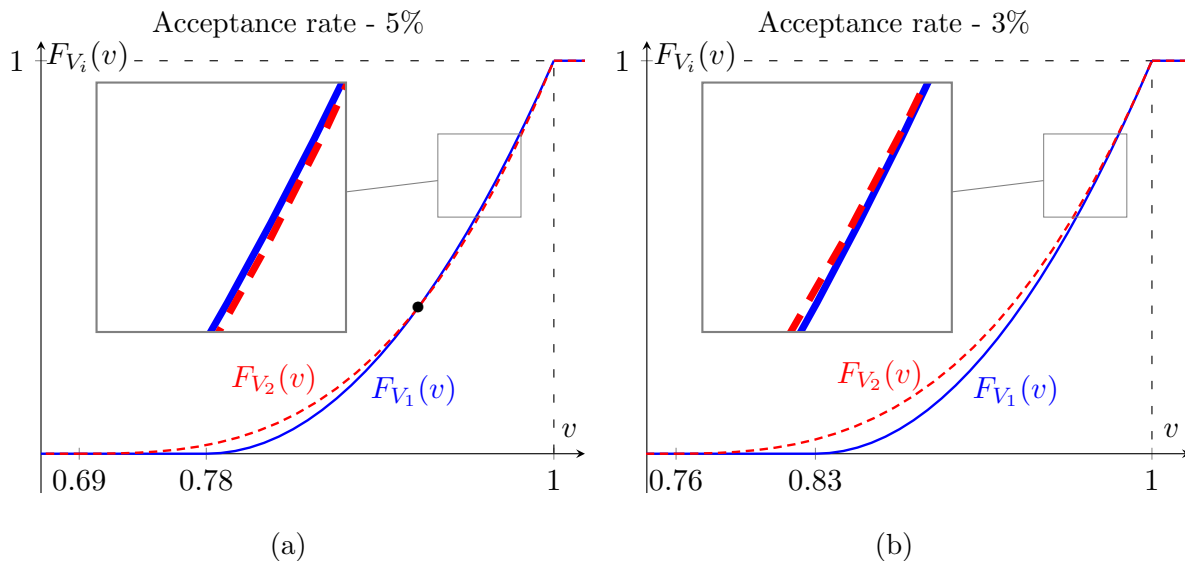


Figure 1: Both graphs relate to the example of one-stage and two-stage screening with $V \sim U[0, 1]$, $N_1 \sim U[-1/4, 1/4]$, and $N_2 \sim U[-1/5, 1/5]$. The solid (blue) lines are the CDFs, denoted $F_{V_1}(v)$, among accepted values under a one-stage screening, while the dashed (red) lines are the CDFs, denoted $F_{V_2}(v)$, among accepted values under a two-stage process. Though first-order stochastic dominance exists under a 3% acceptance rate (or below), but not under a 5% rate (see single crossing at solid black dot), the expected value under one-stage screening is strictly higher in both cases.

expected outcome could go either way.

This result extends beyond dynamic screening. Specifically, one can get similar results when using several evaluators at the same time, rather than one. Take the given example of the TL’s one-stage screening, and consider adding another TL who will conduct an independent screening simultaneously such that acceptance is granted under a unanimity rule. The additional interview is conducted under the same conditions as the original one, namely a noise $N_2 \sim U[-1/4, 1/4]$, and both interviewers are using the same threshold criterion for screening. Under an overall acceptance rate of 5%, the single TL will outperform the unanimous screening of the two, while every acceptance rate below 4.9% dictates a first-order stochastic dominance of the single interviewer over the combination of the two.² The graphs in Figure 2 portrait the two relevant distributions given an overall acceptance rate of 3%.

The next question to address concerns robustness, and specifically, what are the conditions to get such sub-optimal outcomes. So first and foremost, this result is in no way limited to uniform

²If we refer back to the opening paragraph, one can see that any research paper (including this one) is also subject to another form of dynamic screening in the review process when several reviewers assess the paper. Although, technically, it is not “dynamic” in time, a unanimity rule assures that it is essentially a dynamic process.

A one-reviewer versus a two-reviewer screening (acceptance rate - 3%)

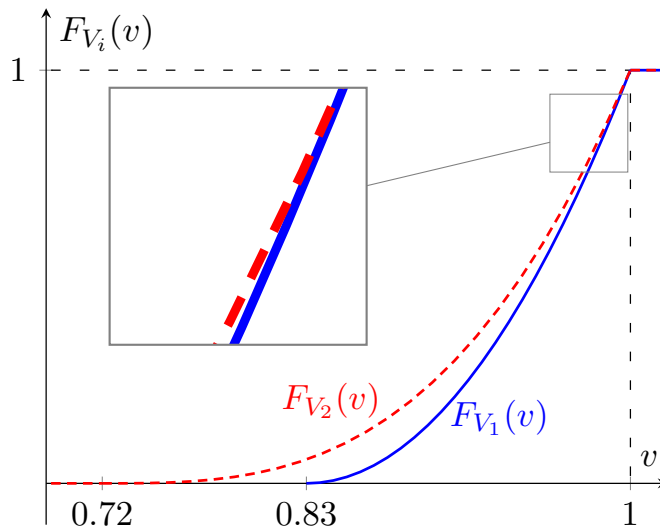


Figure 2: The graphs describe the example of a one-reviewer versus a two-reviewers screening process, where $V \sim U[0, 1]$, and $N_1, N_2 \sim U[-1/4, 1/4]$ under an overall acceptance rate of 3%. The solid (blue) lines represent the CDF, denoted $F_{V_2}(v)$, among accepted values under a two-reviewers screening, while the dashed (red) lines represent the CDF, denoted $F_{V_1}(v)$, among accepted values under a one-reviewer process. Note that the former (first-order) stochastically dominates the latter.

distributions, and we provide similar results for arbitrary bounded non-atomic distributions. Second, as we already know,³ a one-stage screening process based solely on the CTO yields better results than a one-stage screening performed by the TL. Thus, the fact that most of the screening is performed in less accurate stages plays an important role in the process. Moreover, this phenomenon completely reverses when transforming from “elite” screening (i.e., screening at the top of the distribution) to “low-level” screening. In other words, the implicit cost becomes evident when the acceptance rate is low, and subsequent stages are more exposed to relatively high distributive noises. If one should examine the previous examples given a high acceptance rate (for example, more than 70%), then the two-stage screening stochastically dominates the one-stage process.

This distinction brings us to the second part of our analysis, i.e., the asymptotic approach. Since the stakes are high, many cutting-hedge institutions conduct prolonged screening processes with multiple stages. It seems quite implausible that all these institutions are all under-performing. So, for the sake of clarity, we do not claim that. Once multiple stages are introduced, then the overall constraint could be maintained by consistently screening in small portions. That is, every stage can support a high acceptance rate, effectively making it a low-level screening, while the final outcome matches the

³See the notion of a *contraction mapping* in Lagziel and Lehrer (2020).

overall constraint.

In light of this understanding, we reach our second main result which establishes a *convergence to perfect screening*. We prove that, even if all stages are subject to the same noise, a multi-stage asymptotic screening process yields a first-best posterior distribution, as if the screening was preformed with no noise whatsoever. We refer to such an outcome as *perfect screening*, and show that standard stationary strategies, such as a fixed-threshold strategy or a fixed-capacity strategy, lead to a perfect screening outcome. Note that we prove this result even if the same noise is used repeatedly, that is, even if the accuracy along the stages does not improve.

A combination of the results that originate from these two research paths provides a (somewhat striking) conclusion: the quality of a screening process, as a function of the number of stages, is not necessarily monotone. To put it differently, introducing one additional referee or one additional stage may be detrimental, whereas adding multiple ones would probably improve the screening. In what follows, we provide the specific details and conditions to support this statement.

1.2 Relation to literature

The accurate position of this work in the literature is anything but trivial. On the one hand, this work joins the extensive theory of statistical decision making, concerning both the dynamic decision problems [going back to the works of Wald (1939, 1947) and Arrow et al. (1949)] and the comparison of information structure.⁴ On the other hand, this study also concerns the theory of information aggregation, which varies from social learning and information cascade to group decisions and committees of experts.⁵

Though these two branches of the literature are extensive and thorough, our work precisely fits neither. There are three basic elements that distinguish this research from previous ones. First, our formulation of screening problems, in the context of statistical decision theory, naturally combines a capacity constraint which is typically missing from the aforementioned studies [with the exemption of Lagziel and Lehrer (2020)]. This constraint is necessary for our analysis and outcomes. Second, many of the above-mentioned studies extended the seminal work of Condorcet (1785) by introducing costly observations and strategic accumulation of information. Our work is more basic in this sense, because it raises the question of whether another signal is beneficial altogether, irrespective of its price, and independently of the evaluators' preferences (thus information cascades and herding are

⁴See, among others, Milgrom (1981), Quah and Strulovici (2009), Ganuza and Penalva (2010), Chambers and Healy (2011), and more recently Athey and Levin (2018) and Lagziel and Lehrer (2020).

⁵See, e.g., Banerjee (1992), Bikhchandani et al. (1992), Glazer and Rubinstein (1998), Dekel and Piccione (2000), Ottaviani and Sorensen (2001), Levy (2004), Levy (2007), Visser and Swank (2007), Gershkov and Szentes (2009), among many others. For a recent extensive survey, see Bikhchandani et al. (2021).

less relevant in our framework). Third, our framework builds on general, non-atomic distributions where the classification of a signal as either “true” or “false” is irrelevant. In our model, every signal provides more information about the actual value, yet as was exemplified in Section 1.1, the signal might still be detrimental. These stark differences are best exemplified by Bikhchandani et al. (2021), who state that “In purely individual decision making, an extra signal always makes an agent weakly better off” (see Section 2.6 therein). Though this statement is completely true for the relevant models of information aggregation and social learning, one of our main results proves the opposite.

Nevertheless, there are some basic similarities between our study and previous ones. The first, rather basic similarity, is the fact that our asymptotic analysis yields the first-best screening outcome. This is also the key insight of Condorcet (1785), with the obvious distinctions from our model: Condorcet (1785) and subsequent works build on a binary state of the world, a majority rule, and a necessary condition concerning the informativeness of signals, which are irrelevant in our framework. Another similarity, mainly related to the research question, arises from the studies of Ben-Yashar and Paroush (2000) and Berend and Sapir (2005), who prove that random committees of at least three experts outperform a single expert. Yet, the key differences between the models and assumptions completely change the outcome under our analysis.

A different research field where our results may be of some interest is data analysis and data cleansing. In the realm of statistical power analysis, researchers need to maintain a sufficiently large sample size to detect an effect of a given size. In this respect, our capacity constraint fits rather naturally. Our results indicate that reducing the sample size does not necessarily improve the inspection of a given hypothesis. Though a direct application goes well beyond the scope of this work (and quite possibility requires a different, extensive study altogether), we do believe that the given results may be of some importance to this field.

1.3 Structure of the paper

The paper proceeds as follows. In Section 2 we describe the basic model and key definitions. In Section 3 we present the main results, divided into two subsections: in Section 3.1 we compare one-stage and two-stage screening processes, and in Section 3.2 we carry out an asymptotic analysis of screening problems. Concluding remarks, including a discussion about cost functions, are given in Section 4.

2 Preliminaries

Consider a set of elements whose intrinsic values are distributed according to a non-constant random variable V , referred to as an *impact* variable. For every $i \in \mathbb{N}$, let N_i be a random variable which

defines the additive evaluation errors in stage i , and referred to as the *stage- i noise*. We generally assume that all noise variables are unbiased — symmetric about zero and jointly independent of V and of each other. A *capacity* $p \in (0, 1)$ dictates the proportion of accepted elements. That is, the screening is constrained by the requirement to accept a fraction p of the proposed elements.

A k -stage screening problem $\text{SP} = (V, \{N_i\}_{i=1}^k, p)$ consists of an impact variable V , noise variables N_1, \dots, N_k , and a capacity p . The screening problem evolves as follows. Denote $V_1 = V$. In each stage $i \geq 1$, the DM observes $V_i + N_i$ and fixes a threshold $t_i \in \mathbb{R}$, so that

$$V_{i+1} \sim V_i | \{V_i + N_i \geq t_i\}, \quad (1)$$

where

$$\Pr(V_i + N_i \geq t_i) = p_i \quad \text{and} \quad \prod_{i=1}^k p_i = p. \quad (2)$$

In words, in every stage i , the DM observes the noisy valuation $V_i + N_i$ and fixes a screening threshold t_i , so that only the elements whose noisy valuations are at least t_i proceed to the subsequent stage. In every stage i , the DM maintains a capacity of $p_i \in [0, 1]$ to support an overall capacity of $p = \prod_{i=1}^k p_i$. Note that we also allow for *infinite* screening problems, denoted $\text{SP} = (V, \{N_i\}_{i=1}^\infty, p)$, by taking an unbounded number of stages.

Given $\text{SP} = (V, \{N_i\}_{i=1}^k, p)$, a *strategy* τ is a sequence of threshold values $\tau = (t_1, \dots, t_k) \in \mathbb{R}^k$. Let $V_{\text{SP}}(\tau)$ denote the post-screening conditional distribution of the accepted elements' values. That is, $V_{\text{SP}}(\tau) = V_{k+1}$ where V_{k+1} is defined according to Eq. (1), Eq. (2), and τ . The main goal of the DM is to maximize the expected value $\mathbb{E}[V_{\text{SP}}(\cdot)]$ of accepted elements.

For tractability, we make the following assumptions: (i) every variable X (of the above) has a density function f_X and a CDF F_X that are fully supported (namely, f_X is strictly positive) on some bounded interval, denoted $[\underline{X}, \overline{X}]$; and (ii) unless stated otherwise, all noises are i.i.d. random variables. As will later become evident, one can relax these assumptions (e.g., by varying the noises along the stages to be more-informative in the sense of Lehmann (1988), or by using a contracting mapping as in Lagziel and Lehrer (2020); see Section 3.2 therein) and still maintain our key insights. To be clear, we do divert from the second assumption, regarding i.i.d. noises, in the first part of our analysis when we examine a two-stage screening with a more informative second stage, relative to the first stage.

2.1 Stationary strategies

Our analysis is based on threshold strategies, and in the context of dynamic screening, one can also consider two types of stationary strategies: fixed-threshold strategies and fixed-capacity strategies. Formally, fix a k -stage screening problem SP , and consider the following two stationary strategies.

The first strategy, referred to as the *fixed-threshold strategy*, dictates that $\tau = (t, \dots, t)$, which means that all threshold values, throughout the k stages, are identical. By continuity, one can fix the (unique) threshold value t to maintain the capacity constraint p , thus the strategy is well defined. The second strategy, referred to as the *fixed-capacity strategy*, dictates that $p_i = p^{1/k}$ for every stage $1 \leq i \leq k$. That is, all threshold values are fixed so that, in every stage, a fraction $p^{1/k}$ of the elements being evaluated proceed to the subsequent stage. We shall use these strategies in both parts of our analysis.

The first part of our analysis also requires a precise definition concerning stochastic dominance of one screening process over another. The following definition captures this notion.

Definition 1. *For every $i = 1, 2$, consider an i -stage screening problem SP_i with a specific strategy τ_i . We say that one-stage screening under SP_1 given τ_1 stochastically dominates two-stage screening under SP_2 given τ_2 if $V_{SP_1}(\tau_1)$ first-order stochastically dominates $V_{SP_2}(\tau_2)$.*

We should clarify that a one-stage screening problem requires no specification regarding the strategy, since the relevant threshold is well defined and unique. Therefore, henceforth we will not specify the type of the one-stage strategies.

3 Main results

Our analysis consists of two parts, divided into the two following subsections: in Section 3.1 we compare one-stage screening with a two-stage process, and in Section 3.2 we adopt an asymptotic approach by focusing on a dynamic screening while substantially increasing the number of stages.

3.1 A comparison of one-stage and two-stage screening

Our comparison of a one-stage screening and a two-stage process consists of four results. First, in Theorem 1, we consider a two-stage screening performed under a stationary, fixed-threshold strategy where all noises are identically distributed. Under these conditions we prove that for a sufficiently low acceptance rate (i.e., elite screening) one-stage screening stochastically dominates the two-stage process. Next, we extend this result in Proposition 1 and in Proposition 2 by omitting the stationary fixed-threshold condition and by allowing the noises' distributions to vary between the two stages. Lastly, in Theorem 2, we show how one can revert these results by taking a sufficiently high capacity constraint, i.e., through low-level screening.

We start with Theorem 1 which relates to i.i.d. noises along with an implementation of a stationary, fixed-threshold strategy in the two-stage process. In other words, Theorem 1 also captures the alternative interpretation of using two evaluators, instead of one, when both are subjected to the same noise and use the same screening strategy. The theorem states that, for a sufficiently low capacity,

namely a very restrictive screening, it is better to use one stage instead of two. (All proofs are deferred to the Appendix.)

Theorem 1. *For every impact and noise variables V and N , there exists $p_0 > 0$ such that for every capacity $p < p_0$, one-stage screening under (V, N, p) stochastically dominates two-stage screening under $(V, \{N, N\}, p)$ given a stationary fixed-threshold strategy.*

Theorem 1 is based on the understanding that an additional stage, under an overall capacity constraint, must be accompanied with lower thresholds, relative to the one-stage screening. In the second stage this reduction poses a problem since the smaller sample size augments the noise’s effectiveness in distorting the underlying distribution of values.

The next two propositions extend Theorem 1 in two ways. Proposition 1 accounts for non-stationary strategies while maintaining i.i.d. noises, and Proposition 2 extends Proposition 1 by also accounting for different noises. These two extensions require some limitations concerning the two-stage screening strategy. Specifically, by allowing *any* two-stage strategy, one can effectively converge to a one-stage process if one threshold is sufficiently low, which would make it redundant. Moreover, in case the noise in the second stage is a contraction of the noise in the first stage, then we know that doing most of the screening in the second stage will guarantee a better expected outcome compared to a one-stage process. Therefore, we introduce the following definitions of *distinct* two-stage strategies.

Definition 2. *Fix $\varepsilon \in (0, 1/2)$ and a two-stage screening problem SP_2 given a strategy $\tau = (t_1, t_2)$. Note that τ defines two specific capacities (p_1, p_2) . We say that τ is ε -distinct (from a one-stage strategy) if $\max\{p_1, p_2\} < 1 - \varepsilon$. In addition, τ is fully ε -distinct if $\varepsilon < p_2 < 1 - \varepsilon$.*

In simple terms, a two-stage screening strategy is distinct from a one-stage strategy if no stage is redundant in the sense that it screens less than a proportion ε of the elements. Moreover, a fully distinct strategy sustains a constraint on the second stage such that, on the one hand, the second stage is not redundant and, on the other hand, not all the screening occurs in the second stage. As stated, these constraints are necessary, otherwise one could effectively eliminate one stage from the two-stage process and generate results that are at least as good as in one-stage screening.

Though restrictive, there is a simple reasoning as to why the mentioned limitations are rather natural and evident. Generally speaking, the advanced stages in many screening processes are more accurate, therefore more costly. For example, interviews with the CTO or seminars in academic institutions consume a lot of time and effort from busy, time-constrained individuals. It would be very costly to conduct most of the screening in these stages rather than eliminate most of the applicants in preliminary stages., e.g., when screening CVs. So, one must limit the capacity of elements that reach these advanced stages, and Definition 2 provides such a limitation. We discuss this aspect more broadly in Section 4.

The following Proposition 1 shows that even if one should divert from stationary fixed-threshold strategies, the dominance of one-stage screening over the two-stage process remains valid, provided that the strategy is ϵ -distinct and the capacity is sufficiently small.

Proposition 1. *Fix $\epsilon \in (0, 1/2)$. For every impact and noise variables V and N , there exists $p_\epsilon > 0$ such that for every capacity $p < p_\epsilon$, one-stage screening under (V, N, p) stochastically dominates two-stage screening under $(V, \{N, N\}, p)$ given an ϵ -distinct strategy.*

The next proposition shows that one can also extend Proposition 1 to different noises in the two-stage process. For this to hold, we consider a fully distinct strategy such that the second stage does not account for most of the process.

Proposition 2. *Fix $\epsilon \in (0, 1/2)$, an impact variable V , and two noise variables N_1 and N_2 . There exists $p_\epsilon > 0$ such that for every capacity $p < p_\epsilon$, one-stage screening under (V, N_1, p) stochastically dominates two-stage screening under $(V, \{N_1, N_2\}, p)$ given any fully ϵ -distinct strategy.*

Let us emphasize that, in Proposition 2, we do not limit the distribution of N_2 relative to N_1 , other than the general condition of a strictly positive density on some interval. This is more than a mere technicality, since it shows that the additional stage can still distort the screening, even if the added stage is very accurate (i.e., even if the noise is relatively mild). In other words, one does not need to devise esoteric distributions to exemplify our results, but simply extend the screening to the top of the distribution.

On the other hand, if the decision maker wishes to extract only a small portion from the bottom of the distribution, then two-stage screening becomes superior relative to the one-stage process. The intuition is that once the sample size in both stages is significantly large, the noises' influence becomes limited, which makes the additional stage worthwhile. The following theorem is a mirror image of Theorem 1 for cases in which the decision maker focuses on low-level screening (i.e., whenever the capacity is relatively high), showing that two-stage screening dominates the one-stage process.

Theorem 2. *For every impact and noise variables V and N , there exists $p_0 > 0$ such that for every capacity $p > p_0$, two-stage screening under $(V, \{N, N\}, p)$ given a stationary fixed-threshold strategy stochastically dominates one-stage screening under (V, N, p) .*

The combination of Theorem 1 and Theorem 2 implies that the superiority of one screening method over another, varying in the number of stages, greatly depends on the capacity constraint. The transition from a sufficiently low to a sufficiently high capacity, given a stationary fixed-threshold strategy, exemplifies how an additional screening stage changes from a burden to an advantage. A question that remains for future research is whether this transition occurs at a single point such that

one-stage screening is superior below a given capacity and inferior above it, or whether this transition occurs in multiple points.

We now proceed to the second part of our analysis to prove that adding a considerable number of stages strictly improves a screening process.

3.2 Convergence to perfect screening

The results in Section 3.1 may provide the false impression that dynamic screening is inefficient. In this section we prove that this conclusion is false by showing that a multi-stage process eventually yields the first best outcome. Theorem 2 provides some intuition for this, since a sufficiently high capacity implies that additional stages only improve the screening, and that is indeed the case when using multiple stages. We establish this conclusion through two supporting results. The first, in Proposition 3, concerns infinite screening problems and shows that any increasing strategy (i.e., a strategy under which threshold values can only increase along the stages) generates the first-best outcome. The second, in Theorem 3, shows that the two previously mentioned stationary strategies generate posterior distributions that converge, in distribution, to the first best result.

3.2.1 A Perfect screening strategy

In every screening problem, the best the DM can strive for is a screening procedure that yields a result as if there is no noise whatsoever — a result that we refer to as a *perfect screening*. Formally, given a screening problem SP, a strategy τ yields a *perfect screening* if $\mathbb{E}[V_{\text{SP}}(\tau)] = \mathbb{E}[V|V \geq v_p]$, where v_p denotes the p quantile of V (i.e., $\Pr(V \geq v_p) = p$). In other words, a *perfect screening* strategy induces an expected value which is equivalent to a screening without noise, while maintaining the same capacity constraint. Note that a perfect screening of τ , under a capacity of p , entails that both variables, $V_{\text{SP}}(\tau)$ and $V|V \geq v_p$, are identically distributed.

Starting with infinite screening problems, we say that an infinite strategy $\tau = (t_1, t_2, \dots) \in \mathbb{R}^\infty$ is *increasing* if $t_{k+1} \geq t_k$ for every $k \geq 1$. That is, an increasing strategy entails that the screening along the stages becomes stricter. The following proposition shows that, in every infinite screening problem, every increasing strategy produces a perfect screening.

Proposition 3. *In an infinite screening problem, every increasing strategy induces a perfect screening.*

The motivation behind the statement and proof of Proposition 3 originates from the fact that suboptimal elements are discarded with some “patience” on the side of the DM. So even if the noise is rather disruptive for screening, for example, an almost-binary noise with a large variance,⁶ a perfect screening remains feasible since suboptimal elements are slowly screened throughout the stages.

⁶These type of noises (and others) are prone to screening biases; see Lagziel and Lehrer (2019) for more details.

3.2.2 Stationary strategies converge to perfect screening

After establishing that every infinite and increasing strategy yields a perfect screening, one may want to consider a more practical finite set-up. In practice, whether we consider screening job applicants or “cleansing” datasets, decision makers cannot feasibly commit to infinite screening stages, which makes the finite set-up the only practical choice. In such scenarios, a basic question is whether simple finite strategies converge to a perfect screening. In this section we tackle this question and provide two stationary strategies that converge to a perfect screening.

As already defined in Section 2.1, we consider the *fixed-threshold strategy* which maintains the same threshold value throughout the stages, and we consider the *fixed-capacity strategy* which dictates the same capacity in all stages. For each of these strategies we shall prove that the induced distribution converges, in distribution, to a perfect screening outcome. More formally, we say that a stationary strategy τ *converges to a perfect screening* if, for every k -stage screening problem $\text{SP} = (V, \{N\}_{i=1}^k, p)$, we have that $V_{\text{SP}}(\tau) \xrightarrow{d} V|\{V \geq v_p\}$ as $k \rightarrow \infty$. In light of Proposition 3, note that the fixed-capacity strategy does not converge to an infinite increasing strategy. In fact, the limit of the fixed-capacity strategy is not well defined for an infinite screening problem with a capacity constraint of $p \in (0, 1)$. Therefore, its convergence to a perfect screening outcome requires a separate approach and proof, given in the following Theorem 3.

Theorem 3. *The fixed-threshold strategy and the fixed-capacity strategy converge to a perfect screening.*

The use of stationary strategies accommodates a gradual screening process such that only considerably low valuations are eliminated, while most elements proceed to subsequent stages. Though one can devise different non-stationary strategies that converge to a perfect screening, it is clear that not all strategies will do so. For example, consider a screening strategy under which the first-stage threshold is too high, for example $t_1 = v_p + \epsilon$ for some small $\epsilon > 0$, such that the capacity constraint is not violated, yet some values above v_p are partially eliminated in the first stage. In such a case, the thresholds in subsequent stages must be lower, and the resulting posterior distribution would not converge to $V|\{V \geq v_p\}$.

4 Concluding remarks

4.1 The cost of screening

Though our comparison of screening processes does not explicitly incorporate a cost function, we do not remain naive for this consideration. The basic intuition is that additional stages and improved accuracy are more costly, so one should balance the two criteria with the superior expected results. For

example, consider a DM who fixes a screening process given that advanced stages are more accurate (under some metric) than previous ones. A cost-minimization analysis would typically dictate that most of the screening, in terms of capacities, is performed in preliminary stages rather than in advanced ones. In other words, since the screening becomes more costly, the DM limits the mass of elements that reaches advanced stages. Therefore, one can focus on the capacities as an implicit proxy/measure for the needed stages and accuracy. Interestingly, our results in Section 3.1, namely Theorem 1 and Propositions 1 and 2, implicitly contain this feature even without an explicit cost function.

More formally, consider a k -stage screening process with the following, illustrative cost function $\sum_i \alpha_i \log\left(\frac{\prod_{j \leq i-1} p_j}{p}\right)$, where $p_0 = 1$ and $\alpha_{i+1} \geq \alpha_i > 0$ are accuracy indices for stages $i + 1$ and i , respectively (i.e., a higher α indicates a more informative screening under some metric, as in Lehmann (1988) or Lagziel and Lehrer (2020), among others). The intuition behind this function is that the cost of every stage i increases with respect to: (i) the informativeness α_i ; and (ii) the mass of the inspected elements $\prod_{j=0}^{i-1} p_j$. For simplicity, consider $k = 2$ and assume the DM is bounded by a binding budget constraint $C > 0$. So, to meet the condition $\alpha_1 \log\left(\frac{1}{p}\right) + \alpha_2 \log\left(\frac{p_1}{p}\right) \leq C$, the DM must ensure p_1 is sufficiently close to p , which in turn pushes p_2 towards 1. It is straightforward to verify that this constraint becomes even more binding as α_2 increases.

One could also take a different perspective altogether, and consider a set-up where the accuracy of every stage/test, and therefore the cost, depends directly on the induced capacities. For example, consider the extreme case where an examination is performed in stage i under a capacity constraint of $p_i = 1$. What would be the cost in this case? Evidently, the cost should be zero because there is no screening de-facto. Thus, one can consider another illustrative cost function of the form $-\sum_i \log(p_i)$, so that the cost of not performing a screening is zero (that is, $\log(p_i) = 0$ if and only if $p_i = 1$), and it increases as the capacity decreases. Under this example, it is clear that $-\sum_i \log(p_i) = -\log(p)$ and the cost depends only on the overall capacity constraint. In such cases, the DM would be solely concerned with comparing the outcomes of the one-stage and the k -stage screening processes.

4.2 Summary

In this paper we presented an analysis of dynamic screening, showing that the quality of the screening process is not necessarily a monotone function in the number of stages. Specifically, one can add a single screening stage such that the overall quality of the screening process decreases, whereas a few additional stages significantly improve the process. There are a few natural questions that arise from our analysis. First, what exactly occurs when reverting from a single-stage process to a high number of stages? It seems reasonable that first-order stochastic dominance does not revert in a single stage, but that there is a slow transition from one posterior distribution to another. Identifying the posterior

distribution's evolution as a function of the number of stages is an important follow-up question, especially for applicative purposes.

Another rather difficult question to tackle in future research relates to the nature of the noise throughout the process. In our framework, noises are given exogenously, while in practice noises are endogenously determined according to the the capacities and feasibility constraints (technical and monetary). Since the technical complexity of such questions is potentially overwhelming, it might be essential (and even more interesting) to adopt a combined empirical-theoretical perspective, where the assumptions and theoretical analysis are based on actual data.

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A Appendices

A.1 Proof of Theorem 1

Proof. Fix V, N , and some capacity $p \in (0, 1)$. Denote the one-stage and the two-stage screening problems by $SP_i = (V, \{N\}_{k=1}^i, p)$, where $i = 1, 2$ respectively. Denote the screening strategy (and threshold) of SP_1 and SP_2 by $\tau_1 = (t_1)$ and $\tau_2 = (t_2, t_2)$, respectively. Recall that we consider a two-stage screening process with a fixed-threshold stationary strategy. Note that

$$\begin{aligned} F_{V_1}(t) &= \Pr(V \leq t | V + N \geq t_1) \\ &= \frac{1}{p} \Pr(V \leq t, V + N \geq t_1) \\ &= \frac{1}{p} \int_{-\infty}^t f_V(x) \Pr(N \geq t_1 - x) dx. \end{aligned}$$

Thus,

$$\begin{aligned} f_{V_1}(t) &= \frac{1}{p} f_V(t) \Pr(N \geq t_1 - t) \\ &= \frac{1}{p} f_V(t) G(t_1 - t) \\ &= \begin{cases} \frac{1}{p} f_V(t) G(t_1 - t), & \text{for } t_1 - \bar{N} \leq t \leq \bar{V}, \\ 0, & \text{otherwise,} \end{cases} \end{aligned}$$

where $G(x) = \Pr(N \geq x)$ is a differentiable and decreasing function, such that $G(x) = 0$ if $x \geq \bar{N}$, and $G(x) = 1$ if $x \leq \underline{N}$. Similarly,

$$\begin{aligned} f_{V_2}(t) &= \frac{1}{p} f_V(t) G^2(t_2 - t) \\ &= \begin{cases} \frac{1}{p} f_V(t) G^2(t_2 - t), & \text{for } t_2 - \bar{N} \leq t \leq \bar{V}, \\ 0, & \text{otherwise.} \end{cases} \end{aligned}$$

The capacity constraint implies that $t_2 < t_1$, and both values converge to $\bar{V} + \bar{N}$ as $p \rightarrow 0$. So, we will prove that V_1 first-order stochastically dominates (FOSD) V_2 by showing that there exists a range of values close to $\bar{V} + \bar{N}$ (equivalently, a range for p close to zero), such that for every t_1 and every $t_2 < t_1$ in that interval, there exists a unique (interior) point $t \in (t_1 - \bar{N}, \bar{V})$ such that $G(t_1 - t) = G^2(t_2 - t)$, namely a single crossing between the two densities.

To simplify the analysis, define the function $J(x)$ to be the linear extension of G outside the interval

$[\underline{N}, \overline{N}]$. Specifically,

$$J(x) = \begin{cases} G(x), & \text{for } x \in [\underline{N}, \overline{N}], \\ d_{\overline{N}}(x - \overline{N}), & \text{for } x \geq \overline{N}, \\ d_{\underline{N}}(x - \underline{N}) + 1 & \text{for } x \leq \underline{N}, \end{cases}$$

where $d_{\overline{N}} = \lim_{x \rightarrow \overline{N}^-} \frac{G(x) - G(\overline{N})}{x - \overline{N}}$ and $d_{\underline{N}} = \lim_{x \rightarrow \underline{N}^+} \frac{G(x) - G(\underline{N})}{x - \underline{N}}$. Thus, $J'(\overline{N}) < 0$ is well defined and strictly negative, since G is strictly decreasing on $[\underline{N}, \overline{N}]$, by the assumption that all variables have strictly positive densities on their compact supports.

Define the function

$$H(t, t_2) = J(t_1 - t) - J^2(t_2 - t).$$

With a slight abuse of notation, we regard t_2 as a variable, independent of either t_1 , or p . Note that H is continuously differentiable, $H(t_1 - \overline{N}, t_1) = 0$, and $H'_t(t_1 - \overline{N}, t_1) = -J'(\overline{N}) + 2J(\overline{N})J'(\overline{N}) = -J'(\overline{N}) > 0$. Thus, we can use the implicit function theorem to establish that there exist a continuously differentiable function $T(\cdot)$ defined on an open interval I_0 around t_1 , such that $H(T(t_2), t_2) = 0$ for every $t_2 \in I_0$, and $T(t_1) = t_1 - \overline{N}$.

Using this result, let us now prove that there exists a neighbourhood $I_1 = (t_1 - \varepsilon, t_1 + \varepsilon) \subseteq I_0$, where $\varepsilon > 0$, such that for every $t_2 \in I_1$, there exists a unique solution $t \in T(I_1)$ for $H(t, t_2) = 0$. Clearly, a solution exists according to the previous result, so we can assume, to the contrary, that for every interval $I_1 \subseteq I_0$ there exists $t_2 \in I_1$ so that the equation $H(t, t_2) = 0$ has at least two distinct solutions in the interval $T(I_1)$. If that is indeed the case, then by the mean-value theorem, there exists a point $t' \in T(I_1)$ so that $H'_t(t', t_2) = 0$. Since H is continuously differentiable, the equality $H'_t(t', t_2) = 0$ must hold for $\varepsilon \rightarrow 0$, which implies that $t' \rightarrow t_1 - \overline{N}$ and $t_2 \rightarrow t_1$. However, $H'_t(t_1 - \overline{N}, t_1) = -G'(\overline{N}) > 0$ as already stated, and so we get a contradiction.

To conclude, we proved that there exists an open interval $I_1 = (t_1 - \varepsilon, t_1 + \varepsilon)$ such that for every $t_2 \in I_1$ there exists a unique solution $t \in T(I_1)$ for $J(t_1 - t) = J^2(t_2 - t)$, and this solution is continuously differentiable in t_2 . We emphasize that both $J(\cdot)$ and $G(\cdot)$ are independent of t_1 and t_2 , and so is the *length* of I_1 (and I_0).

We now revert to our original notations where t_1 and t_2 sustain any given capacity $p \in (0, 1)$ in the one-stage and two-stage screening processes, respectively. Since $t_2 < t_1$ and both values simultaneously converge to $\overline{V} + \overline{N}$ when $p \rightarrow 0$, one can take $p_0 > 0$ such that $(t_2 - \overline{N}, \overline{V}) \subset I_1$, for every $p \in (0, p_0)$. Moreover, the fact that both densities equal zero at the LHS of their supports (and $\underline{V}_2 < \underline{V}_1$) suggests that $f_{V_2}(t) > f_{V_1}(t)$, in some $[t_2 - \overline{N}, t_p]$, where t_p is the single-crossing point for the chosen capacity. This single-crossing point exists due to our construction, and due to the fact that both screening processes must obey the some capacity constraint (otherwise $f_{V_2}(t) > f_{V_1}(t)$ for every t in both

supports). In other words, for every $t \in [t_2 - \bar{N}, \bar{V}]$, it follows that $G(t_1 - t) \leq G^2(t_2 - t)$ and $f_{V_1}(t) \leq f_{V_2}(t)$, for $t \leq t_p$. This establishes that V_1 (first-order) stochastically dominates V_2 . ■

A.2 Proof of Proposition 1

Proof. This proof follows a construction and arguments similar to those in the proof of Theorem 1. Fix $\epsilon > 0$, random variables (V, N) , and some capacity $p \in (0, \epsilon)$. Denote the one-stage and the two-stage screening problems by $SP_i = (V, \{N\}_{k=1}^i, p)$, where $i = 1, 2$ respectively. Denote the screening strategy (and threshold) of SP_1 and SP_2 by $\tau_1 = (t_1)$ and $\tau_2 = (t_2^1, t_2^2)$, respectively. With no loss of generality (as the stages are interchangeable), we assume that $t_2^1 > t_2^2$. Similarly to the proof of Theorem 1, we have

$$f_{V_1}(t) = \begin{cases} \frac{1}{p} f_V(t) G(t_1 - t), & \text{for } t_1 - \bar{N} \leq t \leq \bar{V}, \\ 0, & \text{otherwise,} \end{cases}$$

where $G(x) = \Pr(N \geq x)$ is a differentiable and decreasing function, such that $G(x) = 0$ if $x \geq \bar{N}$, and $G(x) = 1$ if $x \leq \underline{N}$, and

$$f_{V_2}(t) = \begin{cases} \frac{1}{p} f_V(t) G(t_2^1 - t) G(t_2^2 - t), & \text{for } t_2^1 - \bar{N} \leq t \leq \bar{V}, \\ 0, & \text{otherwise.} \end{cases}$$

The capacity constraint implies that $t_2^1 = \max\{t_2^1, t_2^1\} < t_1$, which suggests that $\underline{V}_2 < \underline{V}_1$ assuming that ϵ and p are small such that $\underline{V}_2 > \underline{V}$ (that is, assuming that the posteriors' lower bounds are strictly above the minimal level \underline{V}). Note that $\underline{V}_1 \rightarrow \bar{V}$ and $\underline{V}_2 \rightarrow \bar{V}$ as $p \rightarrow 0$, whereas $\bar{V}_2 = \bar{V}_1 = \bar{V}$ independently of p . Namely, to reduce the capacity towards zero, the thresholds t_1 and t_2^1 need to increase towards the upper bound of $\bar{V} + \bar{N}$. Therefore, as established in the proof of Theorem 1, the graphs $f_{V_1}(\cdot)$ and $f_{V_2}(\cdot)$ must intersect in at least one interior point, otherwise $f_{V_2}(t) > f_{V_1}(t)$ for every $t \in (t_2^1, \bar{V})$, which violates the capacity constraint for one of the variables.

We will prove that V_1 FOSD V_2 by showing that there exists a range of values for t_1 close to $\bar{V} + \bar{N}$ (equivalently, a range for p close to zero), such that for every $t_2^1 < t_1$ in that interval, there exists exactly one (interior) point $t \in (t_1 - \bar{N}, \bar{V})$ so that $G(t_1 - t) = G(t_2^1 - t)G(t_2^2 - t)$, a single crossing between the two densities.

Define the function $H(t, t_2^1|t_1, t_2^2) = G(t_1 - t) - G(t_2^1 - t)G(t_2^2 - t)$. Again, with some abuse of notation, we allow for t_1 and t_2^1 to vary, independently of p . Assume, to the contrary, that for every interval $I = (\bar{V} + \bar{N} - c, \bar{V} + \bar{N})$ where $c > 0$ and $t_1 \in I$, there exists a $t_2^1 \in (\bar{V} + \bar{N} - c, t_1)$, and a $t_2^2 < t_2^1$ which maintain the ϵ -distinct property, so that the equation $H(t, t_2^1|t_1, t_2^2) = 0$ has at least two

distinct solutions w.r.t. t in the interval $I_1 = (t_1 - \bar{N}, \bar{V})$. The fact that t_2^2 maintains the ε -distinct property suggests that $p_2 = \Pr(X + N > t_2^2) < 1 - \varepsilon$ where $X \sim V|\{V + N \geq t_2^1\}$. Thus, there exists $\delta > 0$ such that

$$t_2^2 > \underline{X} + \underline{N} + \delta = t_2^1 - \bar{N} + \underline{N} + \delta,$$

since $\underline{X} = t_2^1 - \bar{N}$, and

$$G(t_2^2 - t_1 + \bar{N}) < G(t_2^1 - t_1 + \underline{N} + \delta). \quad (3)$$

By the mean-value theorem, there exists a point $t' \in I_1$ such that $H'_t(t', t_2^1|t_1, t_2^2) = 0$. Since H is continuously differentiable, the equality $H'_t(t', t_2^1|t_1, t_2^2) = 0$ must hold for $c \rightarrow 0$, or equivalently, for $t' \rightarrow t_1 - \bar{N}$ and $t_2^1 \rightarrow t_1$. However,

$$\begin{aligned} H'_t(t_1 - \bar{N}, t_1|t_1, t_2^2) &= -G'(\bar{N}) + G'(\bar{N})G(t_2^2 - t_1 + \bar{N}) + G(\bar{N})G'(t_2^2 - t_1 + \bar{N}) \\ &= -G'(\bar{N}) + G'(\bar{N})G(t_2^2 - t_1 + \bar{N}) \\ &> -G'(\bar{N}) + G'(\bar{N})G(t_2^1 - t_1 + \underline{N} + \delta) > 0, \end{aligned}$$

where the first inequality follows from Eq. (3) and the second inequality follows from the facts that t_2^1 is arbitrarily close to t_1 and $G(t_2^1 - t_1 + \underline{N} + \delta) < 1$.

To conclude, we proved that there exists an open interval $I = (\bar{V} + \bar{N} - c, \bar{V} + \bar{N})$, where $c > 0$ and $t_1 \in I$, such that for every $t_2^1 \in (\bar{V} + \bar{N} - c, t_1)$ and $t_2^2 < t_2^1$ which is ε -distinct, there exists a unique solution $t \in (t_1 - \bar{N}, \bar{V})$ such that $G(t_1 - t) = G(t_2^1 - t)G(t_2^2 - t)$. In other words, there exists $p_\varepsilon > 0$ so that for every $p \in (0, p_\varepsilon)$, the densities $f_{V_1}(t)$ and $f_{V_2}(t)$ coincide exactly once on their support (while $\underline{V}_2 < \underline{V}_1$), thus establishing that V_1 (first-order) stochastically dominates V_2 , as needed. ■

A.3 Proof of Proposition 2

Proof. This proof follows the pattern of the proof of Proposition 1, with the additional requirement to show that the lower bound \underline{V}_2 of the conditional impact variable under the two stage screening is indeed lower than the one under the one-stage screening, i.e., $\underline{V}_2 < \underline{V}_1$.

Fix $\varepsilon \in (0, 1)$ and two screening problems, $\text{SP}_1 = (V, \{N_1\}, p)$ and $\text{SP}_2 = (V, \{N_1, N_2\}, p)$, where $p < \varepsilon$. We start by proving that $\underline{V}_2 < \underline{V}_1$. For this purpose consider the following claim

Note that $\tau = (t_2^1, t_2^2)$ is fully ε -distinct, thus $p_2 \in (\varepsilon, 1 - \varepsilon)$ while $p_1 \cdot p_2 = p$. Therefore, if p is sufficiently small (close to zero; specifically $p \ll \varepsilon$), then p_1 tends to 0 and $t_2^1 \rightarrow \bar{V} + \bar{N}$. So, from some point onwards (namely, as long as t_2^1 is sufficiently close to $\bar{V} + \bar{N}$), the support of $V|\{V + N_1 \geq t_2^1\}$ reduces towards \bar{V} , and according to Lemma 1 below, this suggests that t_2^2 is bounded away from $\underline{X} + \bar{N}_1$ where $x \sim V|\{V + N_1 \geq t_2^1\}$. In other words, for every sufficiently small p , the lower bound \underline{V}_2 of V_2 is $t_2^1 - \bar{N}_1$. Recall that $t_2^1 < t_1$ (otherwise, $p_1 < p$), so we conclude that for sufficiently

small capacities $\underline{V}_2 < \underline{V}_1$, as already stated. So, from this point onwards, consider $p < \varepsilon$ such that $|t_2^1 - (\bar{V} + \bar{N}_1)| < \delta$, where $\Pr(N_1 \geq \bar{N}_1 - \delta) < \varepsilon$ for every fully ε -distinct two-stage screening strategy.

A similar computation to the one presented in the proofs of Theorem 1 and Proposition 1 yields

$$f_{V_1}(t) = \begin{cases} \frac{1}{p} f_V(t) G_1(t_1 - t), & \text{for } t_1 - \bar{N}_1 \leq t \leq \bar{V}, \\ 0, & \text{otherwise,} \end{cases}$$

and

$$f_{V_2}(t) = \begin{cases} \frac{1}{p} f_V(t) G_1(t_2^1 - t) G_2(t_2^2 - t), & \text{for } t_2^1 - \bar{N}_1 \leq t \leq \bar{V}, \\ 0, & \text{otherwise,} \end{cases}$$

where $G_i(x) = \Pr(N_i \geq x)$ for $i = 1, 2$, with the same properties as discussed in the previous proofs.

The two screening processes can simultaneously support the capacity p only if the two densities intersect at least once. Let us now show that they do so exactly once. Define the function $H(t, t_2^1|t_1, t_2^2) = G_1(t_1 - t) - G_1(t_2^1 - t)G_2(t_2^2 - t)$. With some abuse of notation, we allow for t_1 and t_2^1 to vary, independently of p , while t_2^2 maintains the fully ε -distinct property.

Assume, to the contrary, that for every interval $I = (\bar{V} + \bar{N}_1 - c, \bar{V} + \bar{N}_1)$ where $c > 0$ and $t_1 \in I$, there exists a $t_2^1 \in (\bar{V} + \bar{N}_1 - c, t_1)$ and a $t_2^2 < t_2^1$ which maintains the fully ε -distinct property, so that the equation $H(t, t_2^1|t_1, t_2^2) = 0$ has at least two distinct solutions w.r.t. t in the interval $I_1 = (t_1 - \bar{N}_1, \bar{V})$. By the mean-value theorem, there exists a point $t' \in I_1$ so that $H'_t(t', t_2^1|t_1, t_2^2) = 0$. Since H is continuously differentiable, the equality $H'_t(t', t_2^1|t_1, t_2^2) = 0$ must hold for $c \rightarrow 0$, or equivalently, for $t' \rightarrow t_1 - \bar{N}_1$ and $t_2^1 \rightarrow t_1$. However,

$$\begin{aligned} H'_t(t_1 - \bar{N}_1, t_1|t_1, t_2^2) &= -G'_1(\bar{N}_1) + G'_1(\bar{N}_1)G_2(t_2^2 - t_1 + \bar{N}_1) + G_1(\bar{N}_1)G'_2(t_2^2 - t_1 + \bar{N}_1) \\ &= -G'_1(\bar{N}_1) + G'_1(\bar{N}_1)G_2(t_2^2 - t_1 + \bar{N}_1) > 0, \end{aligned}$$

where the inequality follows from fact that $\tau = (t_2^1, t_2^2)$ is a fully ε -distinct strategy, and as such $G_2(t_2^2 - t_1 + \bar{N}_1)$ is bounded away from 0 and 1. Hence, $H'_t(\cdot, t_2^1|t_1, t_2^2) > 0$ remains strictly positive as p tends to zero, and there is a single crossing between the two densities f_{V_2} and f_{V_1} . Since $\underline{V}_2 < \underline{V}_1$, we conclude that V_1 FOSD V_2 , as needed. \blacksquare

Lemma 1. *For every noise variable N and every $\varepsilon > 0$, there exists $c > 0$ such that for every impact variable V supported on an interval I of length $c' < c$, the inequality $\Pr(V + N \geq t) > \varepsilon$ implies that $t < \underline{V} + \bar{N}$.*

Proof. Fix a noise variable N . Take $N_0 \in \text{Supp}(N)$ such that $\Pr(N \geq N_0) = \varepsilon$. Fix $c = \bar{N} - N_0$ and fix $a \in \mathbb{R}$. Consider the interval $I = [a, a + \bar{N} - N_0]$ and an impact variable V supported on I .

Take $t = \underline{V} + \bar{N} = a + \bar{N}$ and compute $\Pr(V + N \geq t)$ as follows:

$$\begin{aligned} \Pr(V + N \geq a + \bar{N}) &= \int_a^{a + \bar{N} - N_0} f_V(x) \Pr(N \geq a + \bar{N} - x) dx \\ &< \int_a^{a + \bar{N} - N_0} f_V(x) \Pr(N \geq N_0) dx \\ &= \varepsilon \int_a^{a + \bar{N} - N_0} f_V(x) dx = \varepsilon. \end{aligned}$$

Thus, for every V supported on an interval whose length is $\bar{N} - N_0$, the inequality $\Pr(V + N \geq t) > \varepsilon$ implies that $t < \underline{V} + \bar{N}$. Since the same computation holds for any interval of length smaller than $\bar{N} - N_0$, the statement holds. \blacksquare

A.4 Proof of Theorem 2

Proof. Fix V, N , and a capacity $p > \Pr(V + N > \min\{\bar{V} + \underline{N}, \underline{V} + \bar{N}\})$. For every $i = 1, 2$, denote the i -stage screening problem by $\text{SP}_i = (V, \{N\}_{k=1}^i, p)$, and denote the screening strategy and threshold of SP_1 and SP_2 by $\tau_1 = (t_1)$ and $\tau_2 = (t_2, t_2)$, respectively. Note that for the given capacity (and for higher capacities as well), $t_1 < \min\{\bar{V} + \underline{N}, \underline{V} + \bar{N}\}$. In addition, we know that $t_2 < t_1$ and both thresholds converge to $\underline{V} + \underline{N}$ as $p \rightarrow 1$.

Consider the construction presented in the proof of Theorem 1. Under SP_1 we get

$$f_{V_1}(t) = \begin{cases} \frac{1}{p} f_V(t) G(t_1 - t), & \text{for } \underline{V} \leq t \leq t_1 - \underline{N}, \\ \frac{1}{p} f_V(t), & \text{for } t_1 - \underline{N} \leq t \leq \bar{V}, \\ 0, & \text{otherwise,} \end{cases}$$

where $G(x) = \Pr(N \geq x)$, and under SP_2 we get

$$f_{V_2}(t) = \begin{cases} \frac{1}{p} f_V(t) G^2(t_2 - t), & \text{for } \underline{V} \leq t \leq t_2 - \underline{N}, \\ \frac{1}{p} f_V(t), & \text{for } t_2 - \underline{N} \leq t \leq \bar{V}, \\ 0, & \text{otherwise,} \end{cases}$$

Since $t_2 < t_1$ along with the condition that densities are strictly positive, we can deduce that $f_{V_2}(t) \geq f_{V_1}(t)$ for every $t \in (t_2 - \underline{N}, \bar{V})$, and the inequality is strict for every $t \in (t_2 - \underline{N}, t_1 - \underline{N})$. Thus, we can establish the stochastic dominance of V_2 over V_1 by showing that there is a single crossing between the two densities in the interval $(\underline{V}, t_2 - \underline{N})$.

The fact that at least one crossing exists is trivial due to the capacity constraint. Thus, let us assume, by contradiction, that there is more than one crossing independently of the high capacity p .

That is, we assume that for every $t_1 < \min\{\bar{V} + \underline{N}, \underline{V} + \bar{N}\}$, there exists $t_2 \in (\underline{N} + \underline{V}, t_1)$ such that $H(t|t_1, t_2) = G(t_1 - t) - G^2(t_2 - t) = 0$ has two solutions in the interval $(\underline{V}, t_2 - \underline{N})$.

Similarly to the proof of Theorem 1, by the mean-value theorem, there exists $t' \in (\underline{V}, t_2 - \underline{N})$ such that $H'(t|t_1, t_2) = -G'(t_1 - t) + 2G(t_2 - t)G'(t_2 - t) = 0$. By continuity, this holds for every t_1 and t_2 close to $\underline{V} + \underline{N}$, so we can take the limit $t_i \rightarrow \underline{V} + \underline{N}$ and $t' \rightarrow \underline{V}$, which yields

$$0 = -G'(\underline{N}) + 2G(\underline{N})G'(\underline{N}) = -G'(\underline{N}) + 2G'(\underline{N}) = G'(\underline{N}) < 0;$$

we reached a contradiction.

Therefore, there exists $p > \Pr(V + N > \min\{\bar{V} + \underline{N}, \underline{V} + \bar{N}\})$ such that the densities $f_{V_2}(t)$ and $f_{V_1}(t)$ intersect only once in the interval $(\underline{V}, t_2 - \underline{N})$, and this establishes that V_2 first-order stochastically dominates V_1 , as needed. \blacksquare

A.5 Proof of Proposition 3

Proof. Fix an infinite screening problem $\text{SP} = (V, N, p)$ with an increasing strategy $\tau = (t_1, t_2, \dots)$. The strategy τ maintains, by definition, the stated capacity $p \in (0, 1)$. Therefore, $t_k \leq \bar{V} + \bar{N}$ for every k , otherwise the induced capacity would be zero. So, the sequence (t_1, t_2, \dots) converges to some $t_\infty \in \mathbb{R}$.

Consider the stage- k conditional distribution of V in some $t \in \mathbb{R}$, denoted $F_{V_k}(t)$. Clearly,

$$\begin{aligned} F_{V_k}(t) &= \Pr(V_{k-1} \leq t | V_{k-1} + N_{k-1} \geq t_{k-1}) \\ &= \frac{1}{p_k} \Pr(V_{k-1} \leq t, V_{k-1} + N_{k-1} \geq t_{k-1}) \\ &= \frac{1}{p_k} \int_{-\infty}^t f_{V_{k-1}}(s) \Pr(N_{k-1} \geq t_{k-1} - s) ds \\ &= \frac{1}{p_k} \int_{-\infty}^t f_{V_{k-1}}(s) \Pr(N \geq t_{k-1} - s) ds. \end{aligned}$$

Thus,

$$\begin{aligned} f_{V_k}(t) &= \frac{1}{p_k} f_{V_{k-1}}(t) \Pr(N \geq t_{k-1} - t) \\ &= \frac{1}{p_{k-1} p_k} f_{V_{k-2}}(t) \Pr(N \geq t_{k-2} - t) \Pr(N \geq t_{k-1} - t) \\ &= \frac{1}{\prod_{i=1}^k p_i} f_V(t) \prod_{i=1}^k \Pr(N \geq t_i - t), \end{aligned}$$

and

$$f_{V_{\text{SP}(\tau)}}(t) = \frac{1}{p} f_V(t) \prod_{i=1}^{\infty} \Pr(N \geq t_i - t).$$

Fix $t \in \mathbb{R}$ and recall that $(t_i)_{i \in \mathbb{N}}$ monotonically converges to t_∞ . If $t < t_\infty - \underline{N}$, then there exists N_0 such that $t < t_i - \underline{N}$ for every $i > N_0$, and $\Pr(N \geq t_i - t) < 1$. Hence, $\prod_{i=1}^{\infty} \Pr(N \geq t_i - t) = 0$. Otherwise, $t \geq t_\infty - \underline{N}$, and $t \geq t_i - \underline{N}$ for every $i \in \mathbb{N}$. This implies that $\Pr(N \geq t_i - t) = 1$ for every $i \in \mathbb{N}$, and $\prod_{i=1}^{\infty} \Pr(N \geq t_i - t) = 1$. We thus conclude that

$$f_{V_{\text{SP}(\tau)}}(t) = \begin{cases} 0, & \text{for every } t < t_\infty - \underline{N}, \\ \frac{1}{p}f_V(t), & \text{for every } t \geq t_\infty - \underline{N}. \end{cases}$$

Clearly,

$$\int_{\mathbb{R}} f_{V_{\text{SP}(\tau)}}(t) dt = \int_{t_\infty - \underline{N}}^{\overline{V}} \frac{1}{p}f_V(t) dt = 1,$$

so $t_\infty = v_p + \underline{N}$, and τ induces a perfect screening. ■

A.6 Proof of Theorem 3

Proof. Fix an impact variable V , a noise variable N , and a capacity p . For every $k \in \mathbb{N}$, let $\text{SP}_k = (V, \{N\}_{i=1}^k, p)$ denote a k -stage screening problem, where (V, N, p) are fixed, and consider the fixed-threshold strategy, $\tau = (t^k, \dots, t^k)$. That is, t^k denotes the fixed threshold of τ in SP_k .

We begin by establishing that $\{t^k\}_{k \geq 1}$ is a decreasing sequence that converges to $v_p + \underline{N}$ as $k \rightarrow \infty$. As shown in the proof of Theorem 3, for every $k \in \mathbb{N}$,

$$f_{V_k}(t) = \frac{1}{p}f_V(t) \prod_{i=1}^k \Pr(N \geq t^k - t) = \frac{1}{p}f_V(t) \left[\Pr(N \geq t^k - t) \right]^k.$$

So, if $t^k \leq t^{k+1}$, then

$$f_{V_k}(t) = \frac{1}{p}f_V(t) \left[\Pr(N \geq t^k - t) \right]^k \geq \frac{1}{p}f_V(t) \left[\Pr(N \geq t^{k+1} - t) \right]^{k+1} = f_{V_{k+1}}(t),$$

and the inequality is strict for every t where $f_V(t)\Pr(N \geq t^k - t) > 0$ and $\Pr(N \geq t^k - t) < 1$. Thus we get a contradiction since V_{k+1} is not normalized, and we deduce that $t^{k+1} < t^k$ for every $k \in \mathbb{N}$. In addition, $t^k \geq \underline{V} + \underline{N}$ for every k ; otherwise, there will be no screening whatsoever and the capacity constraint would be violated. Hence, we conclude that $\{t^k\}_{k \geq 1}$ is a decreasing and bounded sequence, and so converges.

To see that $\{t^k\}_{k \geq 1}$ converges to $v_p + \underline{N}$, consider the infinite sequence $(v_p + \underline{N}, v_p + \underline{N}, \dots)$. The induced conditional distribution of V in the respective (infinite) screening problem SP is

$$f_{V_{\text{SP}}}(t) = \frac{1}{p}f_V(t) \prod_{i=1}^{\infty} \Pr(N \geq v_p + \underline{N} - t) = \begin{cases} 0, & \text{for every } t < v_p, \\ \frac{1}{p}f_V(t), & \text{for every } t \geq v_p. \end{cases}$$

If one should take an infinite strategy (t, t, \dots) so that $t \neq v_p + \underline{N}$, the last computation shows that the induced capacity would not be p . So, we conclude that $\lim_{k \rightarrow \infty} t^k = v_p + \underline{N}$.

Finally, we need to prove that V_k converges in distribution to $V|\{V \geq v_p\}$. Fix $t \in \mathbb{R}$. If $t > v_p$, then there exists $k_t \in \mathbb{N}$ such that $t^k - \underline{N} < t$ for every $k \geq k_t$, and $\Pr(N \geq t^k - t) = 1$. Thus, $f_{V_k}(t) = \frac{1}{p} f_V(t)$ for every $t > v_p$ and every $k \geq k_t$. On the other hand, if $t < v_p$, then $\Pr(N \geq t^k - t) < \Pr(N \geq v_p + \underline{N} - t) < 1$ for every $k \in \mathbb{N}$, which yields

$$f_{V_k}(t) = \frac{1}{p} f_V(t) \left[\Pr(N \geq t^k - t) \right]^k < \frac{1}{p} f_V(t) [\Pr(N \geq v_p + \underline{N} - t)]^k \rightarrow 0 \text{ as } k \rightarrow \infty,$$

as needed.

We now move on to the fixed-capacity strategy. Again, for every $k \in \mathbb{N}$, let $\text{SP}_k = (V, \{N\}_{i=1}^k, p)$ denote a k -stage screening problem where (V, N, p) are fixed, and consider the fixed-capacity strategy, $\tau = (t_1^k, t_2^k, \dots, t_k^k)$. One can easily show that, upon a screening stage, the induced posterior distribution of the impact variable first-order stochastically dominates the prior distribution (in fact, the two distributions sustain the monotone likelihood-ratio property). Thus, to maintain a capacity of $p^{1/k}$ in every stage, the thresholds need to strictly increase.

We start by proving that for every $\epsilon > 0$ and every $l \in \mathbb{N}$ there exists $K_{l,\epsilon} \in \mathbb{N}$ such that for every $k > K_{l,\epsilon}$ there are at least l^2 stages where the threshold is strictly higher than $v_p + \underline{N} - \epsilon$, i.e., $|\{i : t_i^k > v_p + \underline{N} - \epsilon\}| \geq l^2$.

Fix $\epsilon > 0$, $l \in \mathbb{N}$, and let $l_k = |\{i : t_i^k > v_p + \underline{N} - \epsilon\}|$ denote the number of stages in which the threshold is above $v_p + \underline{N} - \epsilon$. Note that the capacity up to stage $k - l_k$ is at least $\Pr(V > v_p - \epsilon)$ since values above $v_p - \epsilon$ pass the screening in these stages. Denote $p_\epsilon = \Pr(V > v_p - \epsilon)$ and note that $1 > p_\epsilon > p$. In the remaining l_k stages, the capacity is fixed to $p^{l_k/k}$ according to τ . Thus, when considering the overall capacity p , we get $p > p_\epsilon \cdot p^{l_k/k}$, which translates to $\frac{l_k}{k} > 1 - \frac{\ln(p_\epsilon)}{\ln(p)} > 0$. In other words, there exists $c > 0$, which is independent of k , such that $l_k > ck$. One can now fix $K_{l,\epsilon} = \frac{l^2}{c}$ to get the needed result.

Let us now use this claim to prove that τ converges to a perfect screening. Fix $t < v_p$. Note $\delta = v_p - t$. Fix $l > 0$ so that $\Pr(N \geq \underline{N} + \frac{\delta}{2}) < 1 - \frac{1}{l}$. Take $K_{l,\delta/2}$ so the previous claim holds, and consider $k > K_{l,\delta/2}$. Then,

$$\begin{aligned} f_{V_k}(t) &= \frac{1}{p} f_V(t) \prod_{i=1}^k \Pr(N \geq t^k - t) \\ &\leq \frac{1}{p} f_V(t) \left[\Pr(N \geq v_p + \underline{N} - \frac{\delta}{2} - t) \right]^{l^2} \\ &= \frac{1}{p} f_V(t) \left[\Pr(N \geq v_p + \underline{N} - \frac{\delta}{2} - v_p + \delta) \right]^{l^2} \\ &= \frac{1}{p} f_V(t) \left[\Pr(N \geq v_p + \underline{N} + \frac{\delta}{2}) \right]^{l^2} \\ &= \frac{1}{p} f_V(t) \left[1 - \frac{1}{l} \right]^{l^2} \rightarrow 0 \end{aligned}$$

as $l \rightarrow \infty$. So, for every $t < v_p$, the density $f_{V_k}(t)$ converges to zero, which necessarily leads to a perfect screening. ■