

# Household Risksharing Channels

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## Abstract

This paper aims to fill the gap on the analysis of risksharing channels at the micro level, both within and across households. Using data from the Bank of Italy's Survey on Household Income and Wealth covering the financial crisis, we are able to quantify in a unified and consistent framework several risksharing mechanisms that so far have been documented separately. We find that Italian households were able to smooth on average about 85% of shocks to household head's earnings in both 2008-2010 and 2010-2012. The most important smoothing mechanisms turn out to be self-insurance through saving/dis-saving (40% and 47% in 2008-2010 and 2010-2012, respectively), and within-household risksharing (16% and 14%). Interestingly, risksharing through portfolio diversification and private transfers are rather limited, but the overall percentage of shock absorption occurring through private risksharing channels hovers around four fifths, as opposed to around one fifth of a shock cushioned by taxes and public transfers, excluding pensions. In addition, by exploiting subjective expectations on the following year's household income, we find significant evidence of a lower degree of smoothing of persistent shocks.

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“...[T]he only way to obtain such measures [of income and consumption] is by imposing an accounting framework on the data, and painstakingly constructing estimates from myriad responses to questions about the specific components that contribute to the total.”

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ANGUS DEATON (1997)

## 1. Introduction

Households lie at the center of economic analysis, as they are the core unit of several decision-making processes and perform many economically relevant roles. In fact, there is a large literature focusing on the many roles that households play, both through market transactions (purchases of goods and services, supply of labor and capital services, management of home productions) and via non-market interactions (mutual assistance). Many of these activities are aimed at sharing risk both among household members and across households.

In fact, since Becker’s contributions (1973, 1974)[12, 13], household economics has often stressed the idea that marriage (formal and informal) fosters risk sharing, as transfers between spouses do achieve some smoothing in individual income streams’ variability. Some authors (for example, Chami and Hess, 2005[22]) have gone as far as to suggest that one of the motivations for marriage is to secure some hedging against income risk. Several applied studies (which most frequently employ micro data) provide some support to the idea that marriage achieves a certain amount of risk sharing (as, for example, in the contributions by Rosenzweig and Wolpin, 1985[73], 1994[72]; Rosenzweig, 1988[70]; Rosenzweig and Stark, 1989[71], and others).

There is, however, another subtle way that marriage may influence risk sharing, as it may be the case that more risk sharing comes at the expense of saving, as long as people feel more secure in their spousal agreement (as suggested, for example, by Devereux and Smith 1994[31]). This might decrease the buffer stock from which consumption shocks get smoothed, by the saving/dissaving channel.

As for risksharing across households, suffice it to note that the modern theory of risksharing has been developed centering on the household (or the individual) as its basic decision unit, entering transactions in the market (Arrow 1964[3], Townsend 1994[79]; see Huang and Litzenberger 1988[45] or Deaton 1992[26] for a systematization).

Yet despite the pivotal role that household risksharing plays in basic economic agents’ decisions, very little empirical research has been devoted to the identification and measurement of the mechanisms through which households cope with the risk of income shocks, both between and within them. To be sure, initial empirical tests of risksharing

were carried out at the micro level (Cochrane, 1991[23]; Mace, 1991[57]; Nelson 1994[62]; Hayashi, Altonji and Kotlikoff 1996[43]; Attanasio and Davis 1996[6], Declich and Ventura 2000[28]; Grande and Ventura 2002[39]; Krueger and Perri 2005[53], 2011[54]; Gervais and Klein 2010[38]); however these studies could only test whether the null hypothesis of full risksharing was rejected or not, without being able to identify or measure the economic mechanisms at work. This is all the more unsatisfactory when one considers that theoretical models predicting partial risksharing have been put forward.<sup>1</sup> On the other hand, the macro literature on interregional/international risksharing - whose theoretical underpinning is typically a representative-agent extension of the basic micro framework - has proceeded much further in the empirical analysis of risksharing channels. After the first regression tests of full risksharing (Canova and Ravn 1996)[17], a vast body of literature has developed, starting with Asdrubali, Sørensen and Yosha (1996)[4], henceforth ASY (1996), with the aim of measuring the extent of risksharing channels across countries (or regions) within a unified framework.<sup>2</sup>

A much larger literature on consumption responses to income shocks has focused on the intertemporal (as opposed to cross-sectional) reallocation of resources, under the (often implicit) assumption that the only shock-absorbing mechanism available to households was lending and borrowing in a bonds-only financial market.<sup>3</sup> In sum, as Blundell, Pistaferri, and Preston (2008)[14] point out, beside household saving and borrowing, there is scattered evidence on the role played by various partial insurance mechanisms on household consumption.

This paper aims to fill the gap on the analysis of risksharing channels at the micro level, both within and across households. Using data from the Bank of Italy's Survey on Household Income and Wealth (SHIW) in 2008-2012, we regress consecutive household income measures (from household non-financial income to household income, to household disposable income) on household head's earnings. By doing so, we are able to quantify in a unified and consistent framework risksharing mechanisms that so far have been documented separately. A well-known mechanism is portfolio diversification, which can be implemented through complete markets for contingent claims or appropriate more parsimonious (and realistic) financial structures. Its role has been studied and quantified by Arrow (1964)[3] and Townsend (1994)[79], among others.<sup>4</sup> Another classical risksharing channel consists of fiscal transfer/tax mechanisms. This has been introduced by Sachs and Sala-i-Martin (1992)[74] and von Hagen (1992)[80]. Dynarski and Gruber (1997)[33]

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<sup>1</sup>Incomplete risksharing may arise due to exogenous factors, such as market incompleteness and transaction costs, or endogenous factors, such as limited commitment or enforceability (see Kehoe and Levine 1993[51], further developed by Kocherlakota 1996, Alvarez and Jermann 2000[2], Krueger and Uhlig 2006[52], Krueger and Perri 2011a[55]) and moral hazard.

<sup>2</sup>Tests of risksharing have also used correlation analysis to identify cross-country or cross-regional risksharing. Examples of this strand of the literature include Backus, Kehoe, and Kydland (1992)[9], Pakko (1997)[65], Hess and Shin (1998)[44] and many others.

<sup>3</sup>See the surveys by Deaton (1992) and Jappelli and Pistaferri (2017).

<sup>4</sup>As mentioned above, many seminal studies on risksharing – which explicitly or implicitly only took into account portfolio diversification – aimed at testing full risksharing, without embarking in its quantification.

study the smoothing effect on US household consumption of government transfers (including retirement income) and taxes separately. For Italy, Dedola, Usai and Vannini (1999)[30], Mélitz and Zumer (1999)[60] and Decressin (2002)[29] carry out analyses of public risksharing, but at a macro level. An important - albeit less studied - channel of consumption smoothing is intra-household risksharing, that is the smoothing of the household head’s income shocks through other members’ income changes. Hayashi, Altonji and Kotlikoff (1996)[43] and Dynarski and Gruber (1997) quantify the role of “wife’s earnings”, finding little effects. On the contrary, García-Escribano (2004)[35] models risksharing within families explicitly, obtaining the opposite result. Informal risksharing between households - through private gifts, transfers, aid and services - has been posited by Cox (1987)[24] and extensively studied in developing economies, but rarely quantified in Western countries, at least in the way we do in our empirical analysis. Finally, household self-insurance through asset accumulation and depletion (lending and borrowing in credit markets) has received the most attention, as it stems from the literature on permanent income/life cycle behavior. A related mechanism of self insurance takes place through the (timing of) durable expenditures (see Attanasio 1999[7] for a discussion), and will also be part of our investigation.

While the basic idea of our paper consists in applying the ASY (1996) methodology to households instead of countries, a mere carry-over of the ASY (1996) SUR estimation to a micro setting would be problematic. Indeed, differences exist between macro data on countries and micro data on households, as: i) the former typically include the entire population, while the latter constitute a sample to make inference on, with consequences in terms of selection bias and representativeness; ii) macro data are typically more reliable, both because they originate from official sources and because they benefit from a sort of “washing out” due to aggregation, whereas the latter may be marred by measurement errors, especially in income variation; iii) by definition, at lower levels of aggregation the socio-demographic and economic factors confounding the relation between consumption and income are more numerous than at higher aggregation levels. Specifically, certain individual characteristics – such as age and the presence of children at different stages of the life cycle, plus other possible predictor covariates affecting preference and smoothing capacity – do not even have an obvious homologue at the aggregate level. Moreover, aggregation may get rid of additional factors, such as temporary or sectorial shocks at the household level.<sup>5</sup> Therefore risk sharing mechanisms at lower aggregation levels can be identified only subject to more controls (demographic, geographic, economic, family-linked) than at higher aggregation levels. These difficulties may partly explain the relative scarcity of studies on risksharing channels at the micro level in the last 20 years.<sup>6</sup>

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<sup>5</sup>While the analysis of aggregate data may, under some hypotheses, also disclose relevant microeconomic dynamics, thus making the so called “ecological inference problem” less relevant, this turns out not to be the case in the study of risk-sharing with micro data.

<sup>6</sup>To the best of our knowledge, only a few papers attempted to measure household risksharing channels. Three of them use a mere transposition of ASY, without an explicit derivation from a theoretical model and without controlling for demographic and economic characteristics of the household (Park and Shin (2010)[66], García-Escribano (2004)[35]) or tackling the issue of the endogeneity of the main regressor in

This paper takes on the task of identifying and measuring household risksharing channels, and addresses the issues outlined above in several ways. First, by focusing on the household head’s income, rather than on the household income, we mitigate endogeneity arising from the joint determination of consumption and hours of work (Dynarski and Gruber 1997) or other household-specific unobservable characteristics. Second, by testing regressions with prime-age household heads, we can avoid issues arising from life-cycle/permanent-income intertemporal choices, and focus on cross-sectional (i.e. risksharing) aspects. Third, we address the issue of measurement errors - which is particularly serious in survey microdata - in the main predictor, by using IV estimation.<sup>7</sup> Fourth, by adopting a specification based on household (head)’s earnings as a regressor (instead of aggregate income), we can more easily address the influence of taste shocks on the risksharing metric.<sup>8</sup>

Our reliance on SHIW data presents advantages which have been rarely exploited by the risksharing literature. Indeed, unlike the PSID - which until recently only collected consumption data on food and housing, and not every year - SHIW surveys collect data on all consumption items at a biannual frequency, providing us with a more complete view of total consumption expenditure. In addition, by using true panels of households over couples of consecutive waves and using first differences, we avoid the inefficiencies of unbalanced data plaguing most previous analyses. Furthermore, unlike CEX data, observations on consumption and incomes in SHIW are collected for coincident periods. As Dynarski and Gruber (1997) point out, the availability of US representative consumption data only in the PSID and CEX surveys has forced researchers to merge them with income data at a higher level of aggregation,<sup>9</sup> but the resulting averaging out of individual earnings variation has been detrimental for risksharing estimates, which are based precisely on those variations.<sup>10</sup>

In terms of strategy, our goal is descriptive, in the sense that we aim to establish stylized facts on the degree of household risksharing; but we accomplish that by means of a causal identification, in the sense that we estimate cross-sectional effects of heads earnings growth on consumption growth, controlling for the other intertemporal/life cycle effects,

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the risksharing equations (Balli et al. (2016)[10]). Two others do not adopt an ASY-like methodology: Hayashi, Altonji and Kotlikoff (1996) only deals with two broad channels (risksharing between and within families – not households), does not quantify them (as it only tests for full risksharing) and estimates them separately, with a risk of overlaps. Dynarski and Gruber (1997) measures the extent of risksharing mechanisms in the US, but without embedding them in a unified, internally consistent and theoretically based framework; as a consequence it is not clear that the various mechanisms identified in the analysis are complementary and their measures do not overlap. None of these studies considers all 7 risk sharing channels analyzed in this paper.

<sup>7</sup>See Nelson (1994).

<sup>8</sup>Indeed, as shown by equation (3), household consumption (growth) depends on aggregate income (growth) and taste shock (growth), but not on idiosyncratic variables, such as household (head)’s income. See Sørensen and Yosha (1998).

<sup>9</sup>See for example Attanasio and Davis (1996).

<sup>10</sup>See Gervais and Klein (2010)[38], who show how Dynarski and Gruber’s estimations of household risksharing are downward biased due to the CEX structure.

and strive to purge the earnings variation from endogenous components such as the change in labor supply. Using our framework, we obtain results that can shed light on household risksharing behavior under several dimensions. First, we find that Italian households were able to smooth about 85% of shocks to household head’s earnings in both 2008-2010 and 2010-2012 spells. Second, perhaps surprisingly, the most important smoothing mechanisms turn out to be self-insurance (i.e. (dis-)saving) and within-household risksharing (i.e. income pooling), which were able to absorb as much as 40% of a shock in 2008-2010 and 27% of a shock in 2010-2012, respectively. Informal risk sharing and the capital income channel play a remarkably negligible role, as their small economic significance is accompanied by statistical non-significance; this result is not totally surprising, given the often limited degree of financial depth uncovered in studies on Italian household portfolios as well as the well-known problem of under-reporting of financial assets in the surveys, with the SHIW not being an exception (D’Aurizio et al. 2006)[25].<sup>11</sup>

While private risksharing buffers the big bulk of a shock, public risksharing only cushions about 20% of a shock in both periods, with taxes smoothing more than transfers. Interestingly, our study uncovers a smoothing role for substitution of goods with different durability, at least in the period 2008-2010. This is consistent with other findings in the literature (see, for example, Cerletti and Pijoan-Mas (2012)[21]), highlighting the role that this substitution plays in the transmission of income shocks to nondurable consumption. The paper proceeds as follows. Section 2 develops the methodology to estimate channels of risk sharing within and between households. Section 3 presents the data. Section 4 illustrates the empirical implementation to quantify risksharing channels. Section 5 discusses the empirical results. Section 6 concludes.

## 2. Methodology

### 2.1. Conceptual framework

This section provides the theoretical foundations of the risksharing mechanisms that help smooth household consumption by absorbing shocks to the household heads’ earnings. Consider a stochastic endowment economy, populated by  $J$  infinitely-lived households exhibiting time-separable Von Neumann-Morgenstern (VNM) expected utility functions over a single nondurable consumption good.<sup>12</sup> Uncertainty is represented by a state variable  $s_t$  which summarizes the history up to time  $t$  and the trajectory to infinity and can take on countably many values at any date  $t$ . The Pareto-optimal consumption allocations can be derived by solving the planning problem of maximizing the weighted sum of individual household utilities subject to the feasibility constraint that in each state of nature the sum of household consumptions cannot exceed the sum of all household

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<sup>11</sup>However, financial capital incomes in our dataset exhibit a limited variability, as they are reconstructed as a linear projection of the different assets’ risk classes held by the households.

<sup>12</sup>Generalization to a production economy (Cochrane 1991) and to a multicommodity environment (Hayashi, Altonji and Kotlikoff 1996) is immediate.

endowments. Following standard treatments, such as Cochrane (1991), the first order conditions for all  $s_t$  look like:

$$(\rho^j)^t \lambda^j U_c(C_t^j, \delta_t^j) = \mu_t, j = 1, \dots, J \quad (1)$$

where  $\rho^j$  is household  $j$ 's factor of time preference,  $\lambda^j$  its Pareto weight,  $\delta^j$  its taste shifter and  $\mu_t$  is the Lagrange multiplier associated with the feasibility constraint, divided by the probability of  $s_t$ . The importance of this condition is that it already shows how at the optimum, households' marginal utility is independent of individual household endowments, given aggregate consumption and the Pareto weights. This is true under the assumption, which is standard in the literature, that time and risk preferences are homogeneous across the population. Dividing the expression (1) at two successive dates can get rid of the time-invariant Pareto weight, yielding:

$$\rho^j \frac{U_c(C_{t+1}^j, \delta_{t+1}^j)}{U_c(C_t^j, \delta_t^j)} = \frac{\mu_{t+1}}{\mu_t}, j = 1, \dots, J. \quad (2)$$

The discounted growth of marginal utility is the same across households. The consequences for household consumption growth can be illustrated specifying a CRRA utility function. In this case,

$$\log\left(\frac{C_{t+1}^j}{C_t^j}\right) = \frac{1}{\gamma^j} \left[ \log\left(\frac{\mu_{t+1}}{\mu_t}\right) - \log\left(\frac{b_{t+1}^j}{b_t^j}\right) - \log(\rho^j) \right] \quad (3)$$

where  $\gamma^j$  is household  $j$ 's risk aversion coefficient and  $b_t^j$  is a multiplicative taste shock.<sup>13</sup> The planner's optimal risksharing solution thus prescribes that household consumption growth - net of preference shocks  $[\log(b_{t+1}^j/b_t^j), \gamma^j, \rho^j]$  - must only depend on aggregate consumption growth represented by  $\log(\mu_{t+1}/\mu_t)$ , and must be independent of idiosyncratic household variables, including household (head)'s endowments.<sup>14</sup> Therefore, optimal risk-sharing implies that idiosyncratic shocks are all smoothed out and pooled in the aggregate, regardless of their stochastic process – that is, whether they are transitory or permanent, anticipated or unanticipated, etc.

Equation 3 constitutes the theoretical ground for all the consumption insurance tests which, since the seminal paper by Cochrane (1991), Mace (1991) and Townsend (1994), have been proposed in the literature, and which in a cross sectional setup consist in

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<sup>13</sup>To relate to the previous notation, observe that  $\delta_t^j = [b_t^j \gamma^j]$ .

<sup>14</sup>As shown by Cochrane (1991) this result can be generalized to other utility functions, even non-separable in leisure. More precisely, the utility function may assume any form (provided it is concave and monotonic), may not be time-separable, may not be a VNM function; in addition, arbitrary shocks may be included.

estimating a simple equation of the form:

$$\log\left(\frac{C_{t+1}^j}{C_t^j}\right) = \alpha + \beta x_{t+1}^j + u_{t+1}^j \quad (4)$$

where  $x_{t+1}^j$  is any individual specific variable which, since the contribution by Mace (1991), has normally been represented by an income related variable. As mentioned above, perfect insurance implies  $\beta = 0$  in equation (4).

Moreover, since the contributions by ASY, Dynarski and Gruber (1997), and many others, the magnitude of  $\beta$  has been interpreted as the extent of departure from a situation of perfect insurance, with respect to the shock variable used in equation (4).

## 2.2. Channels of risksharing

The optimal planner solution can be decentralized and implemented through several smoothing mechanisms, depending on the financial and institutional structure of the economy. All these mechanisms provide, in full or in part, a buffer to idiosyncratic shocks, so as to induce a cross-sectional pattern of consumption which is smoother than income. For example, the existence of complete markets of Arrow-Debreu contingent claims (Arrow, 1964), or a specific set of securities (Duffie and Huang, 1985[32]), allows households to implement the full risksharing solution through asset diversification. Similarly, the existence of appropriate government tax/transfers mechanisms allows insuring, at least partially, households whose head's non-financial income has been hit by a negative shock, drawing from incomes hit by a positive shock. In addition, risksharing can be provided through self insurance, that is by asset accumulation (saving) and depletion (dissaving) through lending and borrowing.<sup>15</sup> To be sure, in this case risk sharing (in the sense of cross-sectional smoothing) is a by-product of intertemporal consumption optimization. In fact, in a bonds-only economy, where this intertemporal reallocation is the only feasible risk sharing mechanism, the optimal risk sharing allocation could still be attained, provided all idiosyncratic shocks are temporary (Baxter and Crucini, 1995[11]; Levine and Zame, 2002[56]; Willen, 1999[81]). A peculiar type of (dis-)saving is represented by the timely purchase of durables, which may constitute an additional channel of self insurance (see Cerletti and Pijoan-Mas (2012)[21]). Furthermore, informal risksharing can take place, especially in developing economies, through private gifts, transfers, aid or services. Finally, risksharing can be attained if the household head's income can be pooled with the income of other household members, so as to attain a smoother consumption at the household level.

Unlike some previous work, we maintain a very general setup by not assuming any particular financial or institutional structure for our economy, and let the empirical analysis reveal whether the extent of risksharing in our sample is full, partial or nil, and through

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<sup>15</sup>Self insurance through (dis-)saving aimed at buffering idiosyncratic risk – i.e., precautionary (dis-)saving – should be distinguished from intertemporal trade during the life cycle.



which channels it is attained. We also refrain from modelling endogenous frictions leading to market imperfections (such as limited commitment or enforceability). In fact, the stylized facts and statistical linkages that we uncover will help shed some light precisely on the most appropriate financial and institutional structure or endogenous market imperfections characterizing the Italian economy in the period under exam.

### 2.3. Empirical Model of Risk Sharing Channels

Equation (3) implies that if risk is fully shared through market or non-market institutions, household consumption growth should not respond to idiosyncratic shocks to household head's earnings growth, irrespective of the data generating process governing the latter. As in Attanasio and Davis (1996), Park and Shin (2010) and Dynarski and Gruber (1997), we operationalize this notion by analyzing the regression coefficient of household non-durable consumption growth on the growth in household head's earnings:

$$\log\left(\frac{C_{t+1}^j}{C_t^j}\right) = \alpha + \beta \log\left(\frac{W_{t+1}^j}{W_t^j}\right) + u_{t+1}^j \quad (5)$$

where the disturbance may include a measurement error. Here the intercept captures the effect on consumption variation of aggregate variables, notably aggregate consumption or aggregate income.<sup>16</sup> It is useful to keep in mind that no risksharing implies that the  $\beta$  coefficient be equal to one (i.e. that any idiosyncratic shock is fully transmitted to consumption). On the other hand, if insurance markets and institutions are perfect, then this coefficient should be zero.<sup>17</sup> Intermediate values can then be interpreted as measuring the degree of risksharing. As pointed out by Dynarski and Gruber (1997) and Fafchamps (2011), the  $\beta$  coefficient captures the extent to which the household manages to smooth consumption in the face of shocks to the head's earnings. In other words,

$$1 - \beta = 1 - \frac{\text{Cov}(\Delta \log C^j, \Delta \log W^j)}{\text{Var}(\Delta \log W^j)} \quad (6)$$

is an appropriate measure of the extent of household consumption smoothing via risksharing. The choice of household head's earnings as the shock variable, instead of the more usual household earnings, presents several advantages: it allows more consistency between the regressor and control covariates, reduces endogeneity issues and allows treating other household members' earnings as risk sharing channels.

The main contribution of the risksharing channels methodology consists in a decompo-

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<sup>16</sup>In some specifications of the risksharing model, the term  $\log(\mu_{t+1}/\mu_t)$  is specified as aggregate consumption growth (e.g. Mace 1991), and at times it is added as a regressor to the income growth measure (eg Obstfeld 1994). However, in a cross-section the aggregate term is replaced by the constant term.

<sup>17</sup>See Blundell, Pistaferri, and Preston (2008).

sition of the overall risksharing measure  $1 - \beta$  into the smoothing contributions of the different risksharing mechanisms mentioned above. For every household, we reconstruct the following variables:

- Head's Earnings (household head's wage income + self employed income + pensions):  $W$
- Household Earnings (household members' wage income + self employed income + pensions):  $H$
- Household Income (i.e., Household Earnings + capital income from real estate and financial assets + end-of-service gratuities):  $K$
- Household Gross Income (Household Income + Public transfers received<sup>18</sup>):  $G$
- Household Disposable Income (Household Gross Income - taxes paid<sup>19</sup>):  $T$
- Household Total Disposable Income (Household Disposable Income + inter-and-intra-generational (private) transfers<sup>20</sup>):  $I$
- Household Total Consumption (Household Total Disposable Income - Household Savings):  $E$
- Household Non-durable Consumption: (Household Total Consumption Expenditure - Household Durable Consumption Expenditure):  $C$

The econometric model is based on the idea that, if two successive income measures do not co-move, the smoothing mechanism represented by their difference is at work. For instance, to the extent that  $H$  and  $K$  do not co-move, it means that financial income flows have provided a smoothing effect. By the same token, to the extent that  $G$  and  $T$  do not co-move, it means that taxes have provided further smoothing. Consider the following identity for every household  $j$ :

$$W^j = \frac{W^j}{H^j} \frac{H^j}{K^j} \frac{K^j}{G^j} \frac{G^j}{T^j} \frac{T^j}{I^j} \frac{I^j}{E^j} \frac{E^j}{C^j} C^j. \quad (7)$$

After taking logs and first differences,

$$\Delta w^j = (\Delta w^j - \Delta h^j) + (\Delta h^j - \Delta k^j) + \dots + (\Delta i^j - \Delta e^j) + (\Delta e^j - \Delta c^j) + \Delta c^j \quad (8)$$

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<sup>18</sup>They include unemployment benefits, mobility allowances and various forms of social assistance payments (such as attendance and disability living allowance) which are directly surveyed in the SHIW plus family allowances (ANF) that are simulated (see Appendix A).

<sup>19</sup>A description of the imputation process of gross incomes is given in Appendix A.

<sup>20</sup>These include gifts and transfers from (non-cohabitant) relatives and friends and maintenance payments. Apart from the latter item this variable is conceivable as adding to T informal transfers between households.

where lowercase letters indicate logs.

Multiplying both sides by  $\Delta w^j$  and taking expectations, and then dividing through by  $\text{Var}(\Delta w^j)$ , we obtain a constrained sum of simple regression coefficients:

$$1 = \frac{\text{Cov}(\Delta w^j, \Delta w^j - \Delta h^j)}{\text{Var}(\Delta w^j)} + \dots + \frac{\text{Cov}(\Delta w^j, \Delta e^j - \Delta c^j)}{\text{Var}(\Delta w^j)} + \frac{\text{Cov}(\Delta w^j, \Delta c^j)}{\text{Var}(\Delta w^j)} \quad (9)$$

or

$$1 - \beta = \beta_H + \beta_K + \beta_G + \beta_T + \beta_I + \beta_S + \beta_D. \quad (10)$$

The overall risksharing measure  $1 - \beta$  is decomposed into 7 coefficients. The first coefficient on the RHS -  $\beta_H$  - represents the slope in a regression of  $\Delta w^j - \Delta h^j$  on  $\Delta w^j$ . If a unit positive shock hits the head's earnings,  $\Delta w^j$  will increase by 1 unit; if household earnings  $\Delta h^j$  also increase by 1 unit - that is, if the shock has passed through to household earnings - then  $\beta_H = 0$ , indicating that no intra-household risk sharing has taken place, whereas if household earnings  $\Delta h^j$  stay put - that is, if the shock has not passed through to household earnings - then  $\beta_H = 1$ , indicating that full intra-household risk sharing has taken place. In general,  $\beta_H$  measures the percentage of head earnings changes that is smoothed within the household. By the same token, the second coefficient -  $\beta_K$  - measures the percentage of earnings changes that is further smoothed by capital incomes; the third and the fourth -  $\beta_G$  - and -  $\beta_T$  - measure the further smoothing provided by transfers and taxes, respectively; the fifth -  $\beta_I$  - represents the share that is further smoothed by informal transfers between households; then  $\beta_S$  is the amount of smoothing provided by saving and dis-saving. Finally,  $\beta_D$  represents possible smoothing to non-durable consumption provided by a variation in the timing of durable expenditures.

The next sections will detail the econometric methodology we use to gauge these coefficients as correctly as possible, addressing the estimation issues arising from our setup.

### 3. Data

Our analysis of household risk sharing uses the panel component of biannual data from the Bank of Italy's SHIW, for the periods 2008-2010 and 2010-2012. The main objective of the survey is to study the economic behavior of Italian households, defined as groups of individuals related by blood, marriage or adoption and sharing the same dwelling. The sample size comprises about 8000 households per year selected from population registers and the survey contains a sizable panel component which allows econometricians to estimate target variables' processes and transitions. The head of the household is the person responsible for the household finance, he/she is the main earner in the family and is labelled with an order number equal to one (NORD=1). The longitudinal component allows us to potentially follow over 50% of the households in two spells of twice-repeated

observations.<sup>21</sup> Data collection is entrusted to a specialized company using professional interviewers and CAPI methodology. The survey collects the following information:

- characteristics of the household and of its members (number of income earners, gender, age, education, job status, industry sector, and characteristics of the dwelling);
- income (wage and salaries, income from self-employment, pensions and other financial transfers, income from financial assets and real estate);
- consumption and saving (food consumption, other nondurables, expenses for housing, health, insurance, spending on durable goods, and household saving);
- wealth in terms of real estate, financial assets, liabilities;
- special modules such as capital gains, inheritance, risk aversion, unpaid work, economic mobility, social capital, tax evasion, financial literacy.

From these items, we reconstructed households' balance sheets, income statements, statements of cash flows and consolidated financial statements, along the lines suggested by Samphantharak and Townsend (2006).

Furthermore, since our data do not allow constructing household members' pre-tax incomes, we proceeded to reconstruct pre-tax incomes using an imputation methodology - through *EGaLiTe* tax-benefit microsimulation model - to recover gross figures for basic income and disentangle household allowances from disposable income.<sup>22</sup>

Our variables are measured as reported in section 2.3, and are all in nominal terms.

Table 1 shows some descriptive statistics summarizing the distribution of the key model variables for the two subperiods. In particular, the first biennium of the crisis (2008-2010) is characterized by a very wide distribution for  $(\Delta w^j)$ , with mean equal to 2% and -2.9% in the first and in the second period, respectively. However, in the same spells the growth of nominal nondurable consumption  $(\Delta c^j)$  is higher for the average household (5.4% and 8.4%, respectively), while a contraction in durable consumption is also recorded, in the first of our two periods. This simple comparison of consumption and income growth is suggestive of a rather large decoupling of income and consumption dynamics, which we will indeed find in the regressions' results.<sup>23</sup>

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<sup>21</sup>In the panel component, the sampling procedure is determined in two stages: (i) selection of municipalities (among those sampled in the previous survey); (ii) selection of households to re-interview. This implies that there is a fixed component in the panel (for instance, households interviewed 10 times between 1994 and 2012, or 4 times from 2006 to 2012) and a new component every survey (for instance, households interviewed only in 2012).

<sup>22</sup>See Appendix A

<sup>23</sup>For earlier years, Padula (2004)[63] and Jappelli and Pistaferri (2006)[47], 2010[49],2011[50]) also employ the SHIW data to study the joint dynamics of household income and consumption.

Table 1: Descriptive statistics of the main variables involved in the models estimation

	Years	Mean	Std. Dev.
Earnings growth (head) ( $\Delta w^j$ )	2008-10	2.0%	32%
	2010-12	-2.9%	38%
Earnings plus pensions growth (household) ( $\Delta h^j$ )	2008-10	3.6%	36%
	2010-12	-2.3%	38%
Income growth (household, incl. capital) ( $\Delta k^j$ )	2008-10	4.6%	39%
	2010-12	-1.7%	40%
Gross income growth (hh, incl. public transfers) ( $\Delta g^j$ )	2008-10	4.8%	37%
	2010-12	-0.2%	36%
Disposable income growth (hh, after-tax) ( $\Delta t^j$ )	2008-10	4.1%	34%
	2010-12	-0.5%	33%
Total disp. income growth (hh, incl. priv. transfers) ( $\Delta i^j$ )	2008-10	3.8%	34%
	2010-12	-0.2%	34%
Total consumption growth (hh, excl. saving) ( $\Delta e^j$ )	2008-10	5.4%	44%
	2010-12	7.7%	42%
Non-durable consumption growth (hh, excl. durables) ( $\Delta c^j$ )	2008-10	8.4%	29%
	2010-12	7.9%	30%

Notes: Current prices.  $N_{2008-10} = 1,163$ ;  $N_{2010-12} = 1,138$  Source: Bank of Italy SHIW 2008-10-12. Panel components for consecutive waves. Selection of prime-age households.

## 4. Estimation

At the empirical level, our baseline estimation model implementing the identity (10) above is the following cross-sectional system of linear equations:

$$\begin{aligned}
 \Delta w^j - \Delta h^j &= \nu_H + \beta_H \Delta w^j + \varepsilon_H^j \\
 \Delta h^j - \Delta k^j &= \nu_K + \beta_K \Delta w^j + \varepsilon_K^j \\
 \Delta k^j - \Delta g^j &= \nu_G + \beta_G \Delta w^j + \varepsilon_G^j \\
 \Delta g^j - \Delta t^j &= \nu_T + \beta_T \Delta w^j + \varepsilon_T^j \\
 \Delta t^j - \Delta i^j &= \nu_I + \beta_I \Delta w^j + \varepsilon_I^j \\
 \Delta i^j - \Delta e^j &= \nu_S + \beta_S \Delta w^j + \varepsilon_S^j \\
 \Delta e^j - \Delta c^j &= \nu_D + \beta_D \Delta w^j + \varepsilon_D^j
 \end{aligned} \tag{11}$$

where the  $\nu$  intercepts capture the effect on the dependent variables of aggregate changes. The equation system accounts for the likely cross-equation error correlations, in view of the symmetric structure of our problem. The ordering of the channels in the variance decomposition – and hence of the equations in (11) – stems from the application to the households income statement of the ordering used in OECDs National Accounts and adopted in the literature on risk sharing channels, both macro (e.g., Sørensen and Yosha 1998), and micro (e.g., Park and Shin 2010). This ordering is “natural”, in the sense that certain variables presuppose others. For example, for an economy the construction of GNI implies the existence of GDP, which generates, say, cross-border wage flows; in turn, GDI must build on a measure of gross income like GNI, on which, say, income taxes are derived. Similarly, household financial income builds on wage earnings, disposable income presupposes total income, and so on. Note that we do not posit any causality ordering (as in a recursive VAR, for example), just a logical one.<sup>24</sup> Before estimating the system in 11 we separately estimate the following single equation which, in view of equation (10), is linearly dependent on the others:

$$\Delta c^j = \nu + \beta \Delta w^j + \varepsilon^j. \tag{12}$$

Note that the sum of the  $\beta$  coefficients from equations (11) equals  $1 - \beta$ , that is the coefficient of equation (12). Hence, to estimate the overall degree of risksharing we may as well estimate this coefficient.

Starting from this baseline estimation, we construct augmented estimations to better pinpoint the values of the coefficients in (10) by addressing potential econometric issues plaguing (11) and (12), as described below.

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<sup>24</sup> By this standard, certain channels in our set-up could indeed be switched. For example, there is no stringent reason why public transfers should precede taxes, or private transfers should follow public transfers. Hence, we carried out a robustness test (see discussion in Subsection 5.2 below) whereby we altered the ordering of those channels. As expected, results are essentially unaffected.

*Measurement errors, preference shocks, omitted variables bias, and endogeneity.* Because of the survey characteristics (e.g., response bias), and the imputation exercise we carried out to recover gross incomes, our data – and particularly earnings – may be subject to measurement errors. This problem is only partially mitigated by the accurate surveying methodology applied in sampling SHIW households and by our use of changes in variables. As is well known, such (classical) measurement error boils down to a bias stemming from the earnings variable. Addressing this bias also corrects the inefficiency associated to the coefficient’s standard error.<sup>25</sup>

A second source of bias is the potential correlation between the earnings growth measure and the household preference variation (taste shifter, risk aversion coefficient and rate of time preference) as well as the leisure measure in case of non-separability of the utility function (see Cochrane 1991). The former is partially addressed by adding demographic and household characteristics; the latter is addressed in part by using household head’s earnings as a regressor (as opposed to household income), in part by including a measure of aggregate leisure, which in our cross-sections amounts to adding an intercept in the regressions.

A third, possibly more important source of bias, is the potential endogeneity of hours worked, as they might be driven by the same underlying factors driving growth in consumption.

We focussed on this problem by adopting a new empirical specification, in which we filter out the average effect of endogenously changing labor supply,  $\Delta hrs$ , from the change in income,  $\Delta w$ , and then use the residual as a pure wage shock, as the main regressor in the risk sharing equation. Although this procedure is not equivalent to running an IV estimation for equation (12), it allows us both to address the issue of potential endogeneity in hours worked as a key component in overall head’s earnings, as their effect is removed, and to get an insight as to whether shocks to different components of earnings are associated to different extents of smoothing, which is an interesting research question, in and by itself.

To run this additional empirical analysis, we use information on the variation of head’s hours worked  $\Delta hrs$ , as well as other predictors of the head’s labor income rate of variation, such as the experience and dummies indicating public sector and gender.

Since  $\Delta hrs$  is recovered from self-reported average weekly worked hours and months spent in employment, it may well suffer from rounding and mis-reporting. This makes a direct use of  $\Delta hrs$  less attractive in the ratio  $\Delta w/\Delta hrs$  to calculate the wage variation component of earnings. In fact, since with survey data both  $\Delta w$  and  $\Delta hrs$  may suffer from non-sampling errors, the ratio is likely to suffer from the so-called “division bias.” A viable alternative is a regression of  $\Delta w$  on  $\Delta hrs$  which, however, must take into account the possible residual endogeneity of  $\Delta hrs$  itself due to the (correlated) measurement errors between  $\Delta w$  and  $\Delta hrs$ .<sup>26</sup>

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<sup>25</sup>As the gross incomes are – almost entirely – deterministic functions of net incomes, we do not adjust the earnings standard errors for generated regressor bias.

<sup>26</sup>The bulk of the correlation between the measurement error in the original variable and the instrument will likely disappear with the time differencing we adopt. For example, if a household head systematically

In practice, we proceed as follows: as a first step of our identification strategy, we estimate the following model through a two-stage least squares regression:

$$\begin{aligned}\Delta w^j &= \beta_{w,0} + \beta_{w,1}\Delta hrs^j + \beta'_{w,2}\mathbf{x}^j + u_1^j \\ \Delta hrs^j &= \beta_{h,0} + \beta'_{h,1}\mathbf{z}^j + \beta'_{h,2}\mathbf{x}^j + u_2^j\end{aligned}\tag{13}$$

where we model together the household head's earnings change, and his/her change in hours worked. We are thus purging the change in earnings from the change in hours worked, possibly accounting for endogeneity in the latter by using suitable instruments (the vector  $\mathbf{z}^j$ ): an important worsening in health conditions and first year child rearing (for females), in the first spell, and the occurrence of unemployment and first year child rearing (for females) for the second spell.

For the common instrument (first-year child rearing) the identifying restriction rests on the following argument: during the period of compulsory absence - which in Italy is equal to three months after delivery - the worker is entitled to maternity allowances and/or indemnities which compensate the wage loss. The idea is therefore that in the first year of a child's life, workers (especially women) certainly experience a reduction in hours worked, with a fall in earnings that is less than proportional to the reduction in hours worked. Importantly, this event should have no direct effect on the wage rate, at least in the short run. A similar line of reasoning applies to the case of sickness leave and thus for the specific instrument of the second spell (marked worsening in health conditions). As for the occurrence of unemployment for prime-age males with children, we argue - supported by much of the empirical literature - that this type of worker is characterized by a fairly rigid (full-time) labor supply (Aaberge et al., 1999[1]) and therefore that his unemployment status is unlikely to be an endogenous decision.

Finally, we are using the residual of the income change equation, i.e.  $\widehat{u}_1^j$  - the estimates of  $u_1^j$  - as a valid measure for the pure wage shock to the head's basic income. This variable, more credibly exogenous with respect to  $\Delta c$ , is then used as the main regressor in the place of  $\Delta w$  in the risk sharing regression:

$$\Delta c^j = \nu + \beta\widehat{u}_1^j + \varepsilon^j\tag{14}$$

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underreports her basic income, the effect will wash out when taking first differences (see Dynarski and Gruber 1997).



and in the related equation system:

$$\begin{aligned}
\widehat{u}_1^j - \Delta h^j &= \nu_H + \beta_H \widehat{u}_1^j + \varepsilon_H^j \\
\Delta h^j - \Delta k^j &= \nu_K + \beta_K \widehat{u}_1^j + \varepsilon_K^j \\
\dots\dots \\
\Delta e^j - \Delta c^j &= \nu_D + \beta_D \widehat{u}_1^j + \varepsilon_D^j
\end{aligned} \tag{15}$$

As a last remark, it should be noticed that in order to meaningfully estimate the risk channels equations in (15), we had to generate successive measures of income, as in (7), consistent with rates of change in head's earnings implicit in  $\widehat{u}_1^j$ . However, by a slight abuse of notation and for ease of exposition, we will use the same coefficients to denote the extent of risk sharing (total and for each channel) across equations (11) and (12), and (14) and (15), even though they are attached to different measures of income shocks. In other words,  $\beta$  will represent overall risksharing, regardless of whether we are using  $\Delta w$  or  $\widehat{u}_1^j$  as the shock variable, and  $\beta_H, \dots, \beta_D$  will represent the smoothing obtained by the various channels, in both cases.

***Household characteristics and life-cycle behavior.*** Household-level data are subject to numerous influences, which are typically controlled for by using an additional set of demographic and economic variables, so that equation (14) above becomes:

$$\Delta c^j = \nu + \beta \widehat{u}_1^j + \boldsymbol{\gamma}' \mathbf{y}^j + \varepsilon^j \tag{16}$$

where  $\mathbf{y}^j$  is a vector including standard controls, as suggested in most research on the topic.<sup>27</sup>

Consequently, the equation system in (15) is also estimated using additional covariates in each equation:

$$\begin{aligned}
\widehat{u}_1^j - \Delta h^j &= \nu_H + \beta_H \widehat{u}_1^j + \boldsymbol{\gamma}'_H \mathbf{y}^j + \varepsilon_H^j \\
\Delta h^j - \Delta k^j &= \nu_K + \beta_K \widehat{u}_1^j + \boldsymbol{\gamma}'_K \mathbf{y}^j + \varepsilon_K^j \\
\dots\dots \\
\Delta e^j - \Delta c^j &= \nu_D + \beta_D \widehat{u}_1^j + \boldsymbol{\gamma}'_D \mathbf{y}^j + \varepsilon_D^j
\end{aligned} \tag{17}$$

Two of these controls are of particular interest: a measure of household's net wealth, and the head's expectation for his/her future replacement rate achievable with the public pension, both alone and interacted with the head's wage shock ( $\widehat{u}_1^j$ ). Not only will these variables control for size effects in consumption but, more importantly, they will also ensure that influences on consumption stemming from life-cycle behavior are mitigated.<sup>28</sup> Additional covariates include changes in household components, possibly controlling for

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<sup>27</sup>See Mace (1991) or Dynarski and Gruber (1997).

<sup>28</sup>Controls for demographic and household characteristics also contribute to minimize the effect of life-cycle behavior.

the dynamics in households' economies of scale and for taste shocks due to changes in the household structure, the initial level in the number of earners, head's (quadratic in) age, the presence of children at different stages of the life cycle, head's sector of employment, possible early retirement or unemployment in the arrival year, house ownership as opposed to tenancy, and geographical area.

Note that the  $\beta$  coefficients in regressions (16) and (17) maintain the property of summing up to unity, as in equations (12) and (11). In fact, it is straightforward to show that – since the set of controls is homogeneous across equations – the  $\beta$ 's sum in regressions (16) and (17) corresponds to the sum of the  $\beta$ 's in simple regressions where each variable is replaced by the residual of its projection onto the control vector  $\mathbf{y}^j$ . In other words, we are recasting the variance decomposition in (9) in terms of the “purged” variables.

The models we present are estimated on a restricted sample of households with prime-age household heads (aged 30-55); moreover, we drop households whose head changes in the two year spell. This choice circumscribes the household heads to inelastic labor suppliers, who are less likely to change their working hours in response to consumption changes; it also mitigates concerns related to life-cycle choices, such as moving from student to worker status, or deciding to retire.<sup>29</sup>

***Heteroskedasticity.*** Though heteroskedasticity problems that are common in cross-sectional data are slightly mitigated by our formulation in terms of percentage variations, standard tests still reveal the presence of this problem both in the equation system 11 and in equation 12. To improve inference we estimate the system by a maximum-likelihood conditional mixed-process estimator (CMP), which produces heteroscedasticity-consistent standard errors.

***Nonlinearities.*** An important source of potential bias might be nonlinearities in the determination of consumption, such as the existence of liquidity constraints. As Dynarski and Gruber (1997) point out, consumption changes may not respond to small and frequent variations in the head's earnings, but they may well suffer from large, low-frequency changes (such as an unemployment spell). Hence, our use of variation in hours worked to purge head's labor income may reveal the existence of such liquidity constrained (or simply rule-of-thumb, myopic) behavior. We also try to mitigate issues related to liquidity constraints by focusing on household heads with positive basic income in the start year.

***Attrition.*** We implicitly address issues of attrition - arising from the unavoidable changes of the sample over time (due to births, deaths, marriages, divorces, new sample units arriving, old sample units dropping) - as we have to limit our sample to a balanced panel of households. In fact, we need observing all households for two periods, to be able to compute rates of change in the relevant variables. As for changes within the same household, we control for the initial number and variation of components. Furthermore, we exclude households whose head changed over time.

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<sup>29</sup>Still there is the chance of early retirement and in few cases it is recorded in our estimation data.

**Outliers.** To deal with influential outliers and high leverage data points, particularly relevant in the case of the head’s income variation ( $\Delta w^j$ ), we trim observations from the tails for which the generic value  $x$  is such that  $x < Q(25) - 3IQR$  or  $x > Q(75) + 3IQR$ , where IQR (Interquartile Range) is equal to the difference between the 75<sup>th</sup> and the 25<sup>th</sup> percentiles. More precisely, we remove 147 and 188 observations in the first and in the second spell, respectively. This leaves us with an estimation sample of, respectively, 1,163 and 1,136 observations in the first and in the second spell.

## 5. Results

This section illustrates the results of the implementation of our econometric model, as laid out in Section 4. Table 2 shows, for both the 2008-2010 and the 2010-2012 spells, the OLS and IV estimations of (13), with the change in head’s hours worked ( $\Delta hrs$ ) as the potentially endogenous regressor.

As exclusion restrictions ( $z^j$ ) we use a dummy for the presence of children younger than one year, for female heads (in both spells), a dummy indicating “marked worsening”<sup>30</sup> in health status compared to two years earlier (first spell only), and a dummy for the occurrence of unemployment for male heads with children (second spell only). The difference between the three instruments lies in the fact that while the first, common one, identifies a shock that is likely temporary (mean reverting), the case of a severe deterioration of health is generally associated with permanent innovations (see for example, Jappelli and Pistaferri, 2010[48]). Regarding unemployment, although the related wage shock may persist in the medium run, it should not be permanent. Hence, we are quite confident that the Local Average Treatment Effect (LATE) we are identifying in the first spell relates to a consumption response to a (mainly) temporary earnings shock, while the second could have a larger persistent component.

As expected these indicators show a negative and significant explanatory power on the variation of head’s hours worked, with The F-test statistic on excluded instruments that is well above the conventional threshold of 10 in both first stage equations in both periods, thus ruling out problems of weakness. Since we have two exclusion restrictions for one potential endogenous regressor in each spell, the structural parameters are - technically - over-identified<sup>31</sup> and we can test the instruments’ orthogonality. In the first period the endogeneity tests do not allow us to reject the null hypothesis of regressor exogeneity while it does reject in the second. Finally, the Hansen J test does not allow rejecting the null of instruments’ orthogonality.

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<sup>30</sup>We build this indicator by comparing contemporaneous and lagged scores for self-reported health status (ranging from 1 = very good to 5 = very bad.) This dummy is set equal to one if the head reports a score greater than or equal to 4 while reported a score less than or equal to 2 in the previous survey wave.

<sup>31</sup>For a discussion of identification issues see Section 4

In light of these results we use the OLS estimator for the first spell sub-sample and the IV for the second one to make correct inferences on the residual  $u_1^j$  (proxy of a wage shock) – whose distributions are reported in Table 3 – and we can estimate equations (14) and (16) with robust OLS and the systems (15) and (17) with CMP-SUR, by using head’s wage shock as the main predictor in all equations.

## 5.1. Overall risk sharing

Table 4 illustrates the results for 2008-2010 and 2010-2012 of our baseline specification (12) as in the ASY original set up (columns 1-2), the specification based on wage shocks ( $u_1^j$ ) without additional controls (14) (columns 3-4), and the full specification (16)(columns 5-6) based on wage shocks ( $u_1^j$ ) with additional controls, whose estimated coefficients are not reported for the sake of space, with the exception of the interaction between the expectation for the future replacement rate achievable with the public pension (*reprate*) and household head’s wage shocks (i.e.  $\hat{u}_1 * reprate$ ). Our preferred estimation (Full model OLS in columns 5-6) shows that Italian households were able to smooth around 85% of a wage shock to the household head in both 2008-2010 and 2010-2012, while the above mentioned interaction is not statistically significant at standard significance levels.<sup>32</sup>

That the  $\beta$  estimated coefficients are not very different across the three specifications is noteworthy, as it suggests that, overall, shocks to basic income are smoothed very much like pure wage shocks, and that our basic econometric model is quite robust. Nevertheless, we advocate the use of our augmented model in terms of  $u_1^j$ , as we believe it yields more reliable results, under general conditions.

Despite slight differences between the various specifications, the qualitative conclusions carry over across all estimations: household risksharing in Italy can smooth more than 80% of a primitive source of shocks such as those to the head’s earnings. This result is consistent with most studies on risksharing in Italy, both at the micro and macro level: for example, at the macro level, Scorcu (1997)[75] and Cellini and Scorcu (2002)[20] for 1971-1993, Pellegrini (1997)[67] and Dedola, Usai and Vannini (1999)[30] for 1983-1992, Mélitz and Zumer (1999)[60] for 1984-1992, Gardini, Cavaliere and Fanelli (2005)[36] for 1960-1995, and Cavaliere, Fanelli and Gardini (2006)[19] for 1960-2001 all find a notable and significant degree of smoothing among Italian regions; at the micro level Krueger and Perri (2011)[54] for 1987-2008 reach results on the overall degree of risksharing which are quite close to ours.

## 5.2. Risk sharing channels

How the overall smoothing breaks down across the seven channels of risksharing we have identified is shown in Table 5, which compares the results for 2008-2010 and 2010-2012 of our baseline system equation specification (11) (columns 1-2), the specification based

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<sup>32</sup>In passing, it is worth noticing that partial lack of insurance is only imputable to negative shocks, as the overall insurance coefficient corresponding to positive shocks is always non statistically significant, and ranges from -0.018 to 0.08. Self insurance is always viable with positive shocks.

on wage shocks without additional controls (15) (columns 3-4), and the full specification (17) (columns 5-6) with controls. The table reveals that self-insurance ( $\beta_S$ ) is the most important smoothing mechanism, which is able to absorb 40% of wage shocks in 2008-2010, and around 47% in 2010-2012. Here, the interaction ( $\hat{u}_1 * rebrate$ ) should disentangle the role of life-cycle/pension motives from precautionary savings. Interestingly enough, the elasticity for the interaction is negative and significant (albeit only in the second spell), revealing a lower shock absorption from savings/dis-savings for those households whose head has a higher-than-average expectation for her replacement rate. This is likely evidence that a higher expected (permanent) income in the old age might lower precautionary savings, and thus attenuate the relevance of this smoothing channel. At a macro level Dedola, Usai and Vannini (1999)[30] find somewhat lower but still notable results for credit market insurance in Italy in 1983-1992.

A form of self insurance through the timing of consumption, the adjustment of durable expenditure, seems to achieve a sizable (about 11%) smoothing effect in the first time spell, while it is not statistically significant in the second. A similar effect has been found by Gervais and Klein (2010) in their OLS estimation of CEX data over the 1980-2002 period. Even more to the point, Kruger and Perri (2011) find that in SHIW data from 1987 to 2008 changes in durables are significantly associated with changes in income but are much smaller than the income changes. Also previous findings showing a substitution between durable and nondurable expenditures in periods of crisis (see, among others, McKenzie 2006[59]) are consistent with our results.

Within-household risksharing ( $\beta_H$ ) is also quite large, as it cushions 16% of the shocks in 2008-2010, and about 14% in 2010-2012. This result is in contrast with the findings on the PSID in Hayashi, Altonji and Kotlikoff (1996)[43] and on both the PSID and the CEX in Dynarski and Gruber (1997)[33] - who find nonsignificant effects of non-head income - but parallels the results on the PSID in García-Escribano (2004)[35] - who uses an ASY (1996)-style measure of smoothing. Our result reflects Mocetti, Olivieri and Viviano (2011)'s[61] finding that the effects of the economic crisis on the Italian labour market have been partly absorbed within the households, thanks to i) the greater diffusion of plurinuclear households (the more adults present the lower the risk of joblessness) and ii) the greater propensity to link household formation to employment status. It is also very interesting to compare the estimate in column (6) with that in column (2). As the coefficient's estimate in (6) is about half of the coefficient's estimate in (2), we might conclude that a) there might be an endogeneity bias affecting the choice of hours worked and that b) the added worker effect is associated mainly to a change in head's hours worked. Capital income risksharing ( $\beta_K$ ) does not seem to play any role, as it is neither clearly positive, nor statistically significant. This result is not really surprising, given the often limited degree of financial depth uncovered in studies on Italian household portfolios as well as the well-known problem of under-reporting of financial assets in household surveys, SHIW not being an exception (D'Aurizio et al. 2006).<sup>33</sup> Moreover, our result is consistent with Massa and Simonov (2006)'s[58] finding that Swedish investors do not hedge but in-

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<sup>33</sup>See Guiso and Jappelli (2000)[41].

vest in stocks closely related to their nonfinancial income. Massa and Simonov document that this is directly related to “familiarity”, that is, the tilt to invest in stocks that are geographically and professionally close to the investor or that have been held for a long period.<sup>34</sup>

To these formal channels we can add the informal one - consisting in private transfers between households ( $\beta_I$ ) - that, however, is not particularly sizeable in either spell of the recession and does not exhibit, on average, statistical significance.

While private risksharing channels buffer about two thirds of a shock in both spells, public risksharing only cushions about 20% of a shock in both spells of time, with taxes smoothing more than transfers. However, it is worth noting that pensions are included in the head’s base income and the tax channel excludes risksharing through tax evasion - a phenomenon which is particularly widespread in Italy and which we could not take into account in the reconstruction of basic incomes.<sup>35</sup> At a macro level, in Italy Decressin (2002) finds similar results and Dedola, Usai and Vannini (1999) even higher coefficients for 1983-1992, whereas Mélitz and Zumer (1999) find the public risksharing channel to be insignificant for 1984-1992.

Looking attentively at the systems of risksharing channels provides a deeper insight into which mechanisms underlie the increase in risksharing when hours worked are controlled for; not surprisingly, we find intra-household income, savings and durables, i.e. the channels which are most dependent on the number of hours worked by the household head.

Also the use of a set of controls does make a difference both in the channels’ estimates and in their precision (see e.g. the  $\beta$  coefficients for intra-household risksharing in 2010-2012, for public taxes/transfers and for durables), again corroborating our modelling choice.<sup>36</sup>

### 5.3. Overall risksharing and shock persistence

As already mentioned, the theory only distinguishes between aggregate and idiosyncratic shocks, irrespective of whether they are permanent or transitory, anticipated or unan-

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<sup>34</sup>We must also report a limited variability of financial capital incomes in the survey that are reconstructed as a linear projection of the different risk classes of the assets held by the households.

<sup>35</sup>The biggest discrepancy between our measure of tax risksharing and the actual tax risksharing including tax evasion risksharing arises in the case the interviewed household head lies on the growth of her gross income (to the tax authorities) but not on the growth of her net income (to SHIW interviewers). In this case the tax risksharing that we measure is presumably smaller than the tax risksharing illicitly attained by the household.

<sup>36</sup>As anticipated in Section 4, we carried out a robustness test to check the ordering of the channels in the estimating equations. As there is no stringent reason why public transfers should precede taxes, or private transfers should follow public transfers, we altered the ordering of those channels. As expected, results are essentially unaffected. As a further stress test, we randomly switched the ordering of several channels, irrespective of their logical placement. In this case, it is possible to show that, if the changes in the incomes’ definitions are small, compared to the total, the impact on coefficients will be very small. The largest difference with the standard ordering’s coefficients only attains the second decimal digit when we move a very relevant channel, i.e. intra-household smoothing, from the first to the last place in the decomposition, i.e. just before saving. Even with such an extreme change in the ordering of channels, however, the overall degree of risksharing and the channels’ relative importance are preserved.

anticipated. However, since we look at deviations from the risk sharing allocation, measured consumption may well behave sub-optimally, reacting differently to permanent vs temporary, or anticipated vs unanticipated shocks, presumably depending on market incompleteness or frictions in the attainment of the optimal risksharing allocation. From this viewpoint, a fuller characterization of the deviation from the optimal risk sharing allocation would therefore also be interesting. Since we work essentially with a couple of “two-period” panels, we could not separate the shocks using time-series techniques; instead, we used responses to survey questions conceived to that purpose. While we could not use the observations on subjective expectations about permanent vs temporary shocks, as they contain too many missing data, we exploited information on subjective expectations about following years household income variation surveyed in the 2012 wave of the SHIW<sup>37</sup> and combined them with realized household head’s earnings changes.

More specifically, we split the sample in positive and negative earnings shock realizations. Then, we used the interaction of a positive (negative) expectation for the following year’s household income with a realized positive (negative) head’s wage shock to identify the effect of a shock that is subjectively expected to be non transitory. In other words, we interpreted the coexistence of a negative (positive) shock to current head’s earnings with a negative (positive) expectation for household income for the following year as an indication of “subjective” persistence of the earnings shock. Then, we assessed whether a head’s earnings shock that is combined with an expectation of the same sign has a different impact on current household consumption than an average idiosyncratic shock to heads earnings. The estimation on the overall household sample (Table B.9) shows that, while the positive interaction detects no difference in the earnings change slope, the negative interaction is positive (with a size of 0.29) and significant.

One might observe that there is a discrepancy between the expectation, which is related to household income, and the idiosyncratic shock, which is related to head’s earnings. However, since the head is usually also the survey respondent and his/her labor income often represents the greatest share of household income (precisely, on average, it amounts to 78% in our estimation sub-sample), we are fairly confident that the assumption of a direct link between realized shock and expectation is reasonable. In any case, we also carried out a robustness check, by restricting the sample to single earners households. Estimates are reported in the Appendix B, Table B.8. We believe findings confirm our conclusions. In the first estimation, which guarantees the tightest link between realization and expectation, while the dummies for positive (Posexp) and negative (Negexp) expect-

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<sup>37</sup>The question is: In 12 months, your family’s income will be (distribute 100 points):

- 1 ... much higher than today (at least 10 percent higher)
- 2 ... slightly higher than today (between 2 and 10 percent higher)
- 3 ... substantially unchanged (no more than 2 percent higher or lower)
- 4 ... slightly lower than today (between 2 and 10 percent lower)
- 5 ... much lower than today (at least 10 percent lower)

tations are non-significant, their interactions with the head's wage shock are significant at the 10% level. In particular the estimated coefficient of interaction is negative (-48%) in the positive shock/positive expectation (i.e. positive persistence) sub-sample, while it is positive (+49%) in the negative shock/negative expectation (i.e. negative persistence) sub-sample. This is equivalent to saying that the excess sensitivity of household consumption to an (head's) earnings shock is, in fact, very low for positive shocks perceived as non transitory and significantly amplified in case of negative non transitory shocks. Our interpretation is that households experiencing a persistent positive shock to the head's earnings can afford a fuller smoothing of their consumption, whereas households experiencing a persistent negative shock to the head's earnings will just reduce their consumption, in line with intertemporal consumption optimization.

It should be noted that the identification in the single-earner sub-sample may be affected by the small sample size, therefore the findings should be read with caution and the signs rather than the exact measures should be emphasized.

## 6. Conclusions

The literature has long raised the question of the economic mechanisms underlying the high degree of risksharing often found in micro data. Indeed, while a stream of the literature has always implicitly assumed that risksharing is carried out solely through portfolio diversification, the emergence of the channels literature has shifted the focus towards the diversity of mechanisms implementing (or preventing) the planner allocation. This paper sheds a light on such risksharing mechanisms operating across households. Hence, for example, our results provide a set of possible mechanisms underlying Krueger and Perri (2011)'s findings in SHIW data of a low correlation between labor income and consumption; even more importantly, our methodology can be carried over to other settings to investigate household risksharing in countries where adequate income and consumption data on households are available.

Our finding of a very large role played by intra-household risksharing bears important consequences also for microeconomic modelling. Indeed, as pointed out by Attanasio and Lechene (2002)[8], the pooling of monetary resources is a necessary condition of the unitary model of household behavior. A high degree of intra-household risksharing also brings about macroeconomic consequences: findings for the US by Halla and Scharler (2012)[42] show that marriages do not just improve the allocation of risk at the individual level, but also have implications for the allocation of risk at the more aggregated state level. Finally, in terms of macro modelling, our results show that the bulk of risksharing takes place either through self-insurance or within the household, that is by using the simplest financial tools available to borrow or lend. This suggests that, in modeling consumption in economies like Italy, a bonds-only financial structure might be enough to support the basic patterns of consumption. Further research should be devoted to assess between-households heterogeneity in terms of risksharing capacity along a number of dimensions such as the position of households in the wealth distribution, access to credit, preferences heterogeneity, and more.



Table 2: IV vs. OLS estimation of head's earnings variation and prediction of the wage shock

Dep. variable: $\Delta w$	(2008-10)		(2010-12)	
	(OLS)	(2SLS – IV <sup>i</sup> )	(OLS)	(2SLS – IV <sup>i</sup> )
$\Delta hrs$	0.081*** (0.0096)	0.082*** (0.0211)	0.091*** (0.0105)	0.160*** -0.03
$Experience$	-0.013** (0.0051)	-0.013** (0.0051)	0.003* (0.0014)	0.002* (0.0014)
$Experience^2$	0.000*** (0.0001)	0.000*** (0.0001)	-	-
$Female$	0.036* (0.0196)	-0.036* (0.0197)	-0.042* (0.0229)	-0.051** (0.0237)
$Public\ sec.$	0.058*** (0.0217)	0.058*** (0.0217)	0.04 (0.0279)	0.05 (0.0283)
$Education$	-	-	0.021** -0.01	0.016* -0.01
$Constant$	0.154*** (0.0572)	0.154*** (0.057)	-0.172*** (0.061)	-0.131** (0.065)
$R^2$	0.07	0.07	0.08	0.04
N. of cases	1,163	1,163	1,136	1,136
F-test of excl. instr.	-	148.03	-	61.99
Hansen's J p-value	-	0.32	-	0.28
Endogeneity test p-val	-	0.99	-	0.03

Notes: OLS and IV estimation of model (13) and prediction of  $\hat{u}_1^j$  with estimator suggested by the endogeneity test.

<sup>i</sup> Instrumented:  $\Delta hrs$ .

<sup>i</sup> Excluded instruments:

1. Health conditions marked worsening (dummy) and 2. First year child rearing for females (dummy) (2008-10);

2. Occurrence of unemployment for males with children (dummy) and 2. First year child rearing for females (dummy) (2010-12).

Source: Bank of Italy SHIW 2008-10-12. Panel components for consecutive waves. Selection of prime-age households.

Table 3: Household head's wage shock ( $\hat{u}_1^j$ ) distributions

(2008-10)				(2010-12)			
Percentiles	Smallest			Percentiles	Smallest		
1%	-.844	-1.07168		1%	-1.150	-1.435	
5%	-.560	-1.062535		5%	-.609	-1.372	
10%	-.383	-1.03908	(Obs 1,163)	10%	-.414	-1.280	(Obs 1,136)
25%	-.133	-1.038761		25%	-.159	-1.229	
<b>50%</b>	<b>.0142</b>	<b>Mean</b>	<b>-0.001</b>	<b>50%</b>	<b>.002</b>	<b>Mean</b>	<b>0.001</b>
	Largest	<b>(Std. Dev. .308)</b>			Largest	<b>(Std. Dev. .369)</b>	
75%	.155	1.070		75%	.177	1.248	
90%	.342	1.079	Var. .095	90%	.392	1.280	Var. .136
95%	.504	1.081	Skew. -.161	95%	.606	1.301	Skew. .035
99%	.897	1.217	Kurt. 4.78	99%	1.181	1.343	Kurt. 5.30

Note: prediction of  $\hat{u}_1^j$  from equation (13).

Source: Bank of Italy SHIW 2008-10-12. Panel components for consecutive waves. Selection of prime-age households.

Table 4: Overall risksharing

	(1) Benchmark OLS		(2) Adjusted OLS		(3) Full model OLS	
Dep. var.: $\Delta c^j$	2008-10	2010-12	2008-10	2010-12	2008-10	2010-12
<i>Unsmoothed consumption</i>						
$[\beta]: (\Delta w^j   \hat{u}_1)$	0.170*** (0.0452)	0.151*** (0.0306)	0.143*** (0.0446)	0.133*** (0.0322)	0.146*** (0.0430)	0.149*** (0.0323)
$[\Delta\beta]: (\hat{u}_1 * reprice)$					-0.066 0.2055)	0.157 (0.1343)
<i>Constant</i>	0.081*** (0.0128)	0.084*** (0.0120)	0.084*** (0.0127)	0.080*** (0.0121)	0.606 (0.4921)	-0.228 (0.5317)
Controls	No	No	No	No	Yes	Yes
$R^2$	0.036	0.037	0.024	0.028	0.120	0.089
N. of cases	1,163	1,138	1,163	1,129	1,163	1,129

Note: Estimation of equations (12) (column 1-2), (14) (column 3-4) and (16) (column 5-6).  $[\beta]$  is the wage shock (i.e.  $\Delta w^j$  in col.1,  $\hat{u}_1$  in cols. 2 and 3) estimated coefficient;  $[\Delta\beta] = \hat{u}_1 * reprice$  is the variation of  $[\beta]$  related to heterogeneity in pension expectations.

Source: Bank of Italy SHIW 2008-10-12. Panel components for consecutive waves. Selection of prime-age households.

Table 5: Risksharing channels

<i>Channels</i>	(1)		(2)		(3)	
	Benchmark system		Adjusted system		Full model system	
	2008-10	2010-12	2008-10	2010-12	2008-10	2010-12
<u>1. Basic income from other members</u>						
$[\beta_H]$	0.159***	0.285***	0.158***	0.173***	0.162***	0.137***
	(0.0406)	(0.0450)	(0.0447)	(0.0515)	(0.0434)	(0.0467)
$[\Delta\beta_H]:(\hat{u}_1 * reprice)$					-0.117	-0.137
					(0.2573)	(0.2129)
<u>2. Capital incomes (financial and real)</u>						
$[\beta_K]$	0.003	-0.023	0.003	0.000	-0.003	-0.003
	(0.0196)	(0.0194)	(0.0198)	(0.0131)	(0.0165)	(0.0113)
$[\Delta\beta_K]:(\hat{u}_1 * reprice)$					-0.023	0.011
					(0.0812)	(0.0525)
<u>3. Public transfers other than pensions</u>						
$[\beta_G]$	0.067***	0.110***	0.064***	0.081**	0.078***	0.079**
	(0.0153)	(0.0268)	(0.0170)	(0.0212)	(0.0132)	(0.0197)
$[\Delta\beta_G]:(\hat{u}_1 * reprice)$					0.224***	0.120*
					(0.1040)	(0.0723)
<u>4. PIT &amp; Property tax on OODs</u>						
$[\beta_T]$	0.106***	0.100***	0.107***	0.092***	0.114***	0.104***
	(0.0106)	(0.0066)	(0.0113)	(0.0107)	(0.0123)	(0.0098)
$[\Delta\beta_T]:(\hat{u}_1 * reprice)$					0.146***	0.152***
					(0.0505)	(0.0457)
<u>5. Informal transfers</u>						
$[\beta_I]$	0.024	-0.001	0.010	0.006	-0.004	0.007
	(0.0173)	(0.0062)	(0.0079)	(0.0056)	(0.0061)	(0.0085)
$[\Delta\beta_I]:(\hat{u}_1 * reprice)$					-0.182**	-0.011
					(0.0866)	(0.0203)
<u>6. Saving/dissaving</u>						
$[\beta_S]$	0.397***	0.313***	0.412***	0.478***	0.401***	0.474***
	(0.0716)	(0.0562)	(0.0714)	(0.0661)	(0.0736)	(0.0594)
$[\Delta\beta_S]:(\hat{u}_1 * reprice)$					-0.118	-0.717***
					(0.3810)	(0.2666)
<u>7. Durable expenditures</u>						
$[\beta_D]$	0.073	0.064*	0.102*	0.038	0.105*	0.054
	(0.0570)	(0.0348)	(0.0537)	(0.0375)	(0.0548)	(0.0343)
$[\Delta\beta_D]:(\hat{u}_1 * reprice)$					0.134	0.424**
					(0.2095)	(0.1678)
	Controls	No	No	No	No	Yes
	N. of cases	1,163	1,136	1,163	1,129	1,163
						Yes
						1,129

Note: Estimation of equation systems (11) (column 1-2), (15) (column 3-4) and (17) (column 5-6):

$\beta_H$  is the head wage shock's estimated coefficient on the dependent var.  $(\Delta w^j | \hat{u}_1) - \Delta h^j$ ;

$\beta_K$  " "  $\Delta h^j - \Delta k^j$ ;

$\beta_G$  " "  $\Delta k^j - \Delta g^j$ ;

$\beta_T$  " "  $\Delta g^j - \Delta t^j$ ;

$\beta_I$  " "  $\Delta t^j - \Delta i^j$ ;

$\beta_S$  " "  $\Delta i^j - \Delta e^j$ ;

$\beta_D$  " "  $\Delta e^j - \Delta c^j$ .

Source: Bank of Italy SHIW 2008-10-12. Panel components for consecutive waves. Selection of prime-age households.

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## Appendix A. Simulation of gross incomes

Income variables in household survey data are often recorded net of income taxes and other levies on income, such as social contributions. However, for many research tasks gross income information is crucial. Examples are the calculation of tax wedges and effective tax rates or issues related to the distribution or determinants of market incomes. Another application of household micro-data where the lack of gross incomes can be a major problem is tax-benefit microsimulation. These models feature detailed social and fiscal policy rules as they apply to individuals and households and are largely used by governments as well as academic researchers. In addition to their main use as tools to analyze the effects of fiscal and social policy measures, these models are used to impute tax figures that are not gathered in the survey questionnaire.

In the case of our analysis, raw data from the survey must be appropriately treated in order to determine the net income for personal income tax (PIT) purposes, then the net-to-gross income procedure can be carried over. Rather than approximating the tax system using a functional form (Blundell *et al.* 2016[15], for instance, use the functional form suggested by Heathcote, Storesletten, and Violante (2014)), we decided to replicate faithfully the Italian taxation system in force at the time using microsimulation.

*EGaLiTe* ([37]) is a static tax&benefit microsimulation model. It uses a standard iterative method to simulate net-to-gross personal income trajectories. The codes are written in STATA. The fiscal module is based on microdata on Italian families provided by the Bank of Italy Survey of Households Income and Wealth (SHIW) that surveys after-tax income variables. The model aims to simulate personal income tax paid by Italian taxpayers (IRPEF) in order to determine the status quo distribution of the tax burden as well as the distributive effects of alternative reforms. In particular it simulates the IRPEF progressive structure, including its regional/local surtaxes and the main tax expenditures. Moreover, it approximates the distribution of family allowances (Assegno al Nucleo Familiare) which represent the main subsidy for households with dependent children in Italy but - unfortunately - cannot be directly disentangled from the labor income information reported in the survey. Finally, the fiscal module simulates the tax impact of owner-occupied dwellings (whose imputed rent is fully deductible from the PIT tax base in the period 2008-2010) which in the second spell is embodied in the new property tax “IMU”. This latter tax-payment for 2012 is self-reported by respondents in the survey.

Since a micro analysis of tax evasion behavior is beyond the scope of this study, we adopt the simplifying assumption of no tax evasion in earnings. This can be easily accepted for employees while bringing lower accuracy in reconstructing gross figures for the self-employed. The loss of accuracy is however mitigated by the fact that we work with changes in variables, and tax evasion in Italy does not tend to vary much over time.

Given the impossibility of analytically deriving an individual measure of gross income starting from net income, an iterative algorithm is adopted (see Sutherland, 2001[78]; Immervoll and O’Donoghue, 2001[46]). In practice it consists in estimating a plausible individual gross value starting from the self-reported disposable amount. Then, the tax rules for obtaining the net value are applied to this gross value. This value is compared with the sample original value and if these are equal - net of a margin of tolerance - the

gross income estimate is considered a good approximation of the unknown value. Outside the tolerance margin, the algorithm predicts a new gross value (larger or smaller, depending on the sign of the error) and applies the tax rules again. This iteration continues until convergence is achieved for all tax payers in the sample. In fact, given the self-reported after-tax income, the characteristics of the tax payer (number of children, dependent spouse, presence of owned properties, mortgages) as well as the potential tax relief for income source plus other allowances and tax expenditures, there is only one taxable income such that, by applying the tax rules, one obtains the original after-tax income.

To determine the tax structure the following steps are followed:

1. identification of total income, i.e. the sum of the different sources of income subject to the IRPEF;
2. simulation and subtraction of the standard deductions (e.g. deduction for owner-occupied housing) from 1. = taxable income;
3. Application of the tax scale (A.6) to 2. = gross tax;
4. Subtraction of income tax credits, relief and tax expenditures from 3. (see A.7 for the employee tax relief pattern) = net tax.

Table A.6: Statutory tax rates and brackets (2008-2012)

<b>Bracket</b> (Euros x 1000)	<b>Tax rate</b> (%)
Up to 15	23
From 15 to 28	27
From 28 to 55	38
From 55 to 75	41
Over 75	43

Table A.7: Employee tax relief (2008-2012)

<b>Total annual income (Y)</b> (Euros x 1000)	<b>Annual deduction</b> (Euros x 1000)
Up to 8	1.840
From 8 to 15	$1.338 + [0.502 * (15 - Y) / 7]$
From 15 to 55	$1.338 * [(55 - Y) / 40]$
Over 55	0

## Appendix B. Expectations and persistent income shocks

Table B.8: Overall risksharing, single earner households

Dep. var.: $\Delta c^j$	(1) (Positive persistent)	(2) (Negative persistent)
$[\beta]: \hat{u}_1$	0.198* (0.1105)	0.137 (0.1167)
Posexp	-0.128 (0.0929)	- -
Negexp	- -	-0.065 (0.0583)
$[\Delta\beta]: \hat{u}_1 * Posexp$	-0.477* (0.2626)	- -
$[\Delta\beta]: \hat{u}_1 * Negexp$	- -	0.492* (0.2510)
Constant	-0.483 (1.0760)	0.153 (1.0817)
Controls	Yes	Yes
$R^2$	0.275	0.188
N. of cases	171	202

Note: Estimation of equations (16) breaking down the sample by the sign of the shock. The interaction of a positive (negative) expectation for the following year's household income with an experienced positive (negative) wage shock is used to identify the differential effect of a "persistent" shock.

$[\beta]$  = wage shock ( $\hat{u}_1$ ) estimated coefficient; Posexp=positive subjective expectation for next year's household income; Negexp=negative subjective expectation for next year's household income;  $\hat{u}_1 * Posexp$  and  $\hat{u}_1 * Negexp$  = shock persistency interactions.

Source: Bank of Italy SHIW 2010-12 panel components. Selection of prime-age, single earner households.

Table B.9: Overall risksharing, all households

Dep. var.: $\Delta c^j$	(1) (Positive persistent)	(2) (Negative persistent)
$[\beta]: \hat{u}_1$	0.045 (0.0613)	0.094 (0.0638)
Posexp	-0.113** (0.0535)	-
Negexp	-	-0.008 (0.0460)
$[\Delta\beta]: \hat{u}_1 * Posexp$	0.115 (0.1687)	-
$[\Delta\beta]: \hat{u}_1 * Negexp$	-	0.287** (0.1412)
Constant	-0.216 (0.6758)	-0.568 (0.7815)
Controls	Yes	Yes
$R^2$	0.132	0.148
N. of cases	578	558

Note: Estimation of equations (16) breaking down the sample by the sign of the shock. The interaction of a positive (negative) expectation for the following year's household income with an experienced positive (negative) wage shock is used to identify the differential effect of a "persistent" shock.

$[\beta]$  = wage shock ( $\hat{u}_1$ ) estimated coefficient; Posexp=positive subjective expectation for next year's household income; Negexp=negative subjective expectation for next year's household income;  $\hat{u}_1 * Posexp$  and  $\hat{u}_1 * Negexp$  = shock persistency interactions.

Source: Bank of Italy SHIW 2010-12 panel components. Prime-age households.