

Radar Clutter Mitigation via Probability Measure Transform

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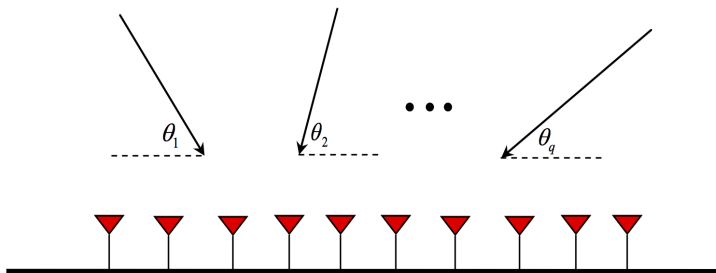
Outline

- ▶ Background and motivation
- ▶ Probability measure transform
- ▶ Measure transformed Gaussian QML estimator
- ▶ Application: robust direction finding in heavy-tailed clutter
- ▶ Summary

Background and motivation

Problem

In practical direction finding of radar targets **accurate parametric models** for the signals and clutter are **unavailable**



Background and motivation

Goal

Estimate $\theta_0 \in \mathbb{R}^m$ given samples $\{\mathbf{X}_n\}_{n=1}^N$ from $P_{\mathbf{X};\theta_0}$

Restriction

$P_{\mathbf{X};\theta_0}$ belongs to *unknown* parametric family of probability measures

$$P_{\mathbf{X};\theta_0} \in \{P_{\mathbf{X};\theta} : \theta \in \Theta\}$$

- ▶ The maximum likelihood estimator cannot be implemented
- ▶ Resort to methods that require *partial statistical information*

Background and motivation

Gaussian QML estimator [White 1982]

- ▶ Minimize the *empirical KLD* between $P_{\mathbf{X};\theta_0}$ and a Gaussian measure $\Phi_{\mathbf{X};\theta}$ with mean $\boldsymbol{\mu}_{\mathbf{X}}(\boldsymbol{\theta})$ and covariance $\boldsymbol{\Sigma}_{\mathbf{X}}(\boldsymbol{\theta})$
- ▶ Amounts to maximization of

$$J(\boldsymbol{\theta}) \triangleq -D_{\text{LD}} \left[\hat{\boldsymbol{\Sigma}}_{\mathbf{X}} \parallel \boldsymbol{\Sigma}_{\mathbf{X}}(\boldsymbol{\theta}) \right] - \|\hat{\boldsymbol{\mu}}_{\mathbf{X}} - \boldsymbol{\mu}_{\mathbf{X}}(\boldsymbol{\theta})\|_{(\boldsymbol{\Sigma}_{\mathbf{X}}(\boldsymbol{\theta}))^{-1}}^2$$

- ▶ The GQMLE:

$$\hat{\boldsymbol{\theta}} = \arg \max_{\boldsymbol{\theta} \in \Theta} J(\boldsymbol{\theta})$$

- ▶ Simple implementation, easy performance analysis
- ▶ Sensitive to *model mismatch* (e.g. in non-Gaussian clutter)

Background and motivation

Proposed approach

GQMLE under a **transformed probability distribution** of the data

Advantages

- ▶ Resilient to outliers
- ▶ Involves higher-order statistical moments
- ▶ Significant **mitigation of the model mismatch effect**
- ▶ Computational advantages of the first and second-order methods of moments

Probability Measure Transform

Definition

Given a non-negative function $u : \mathcal{X} \rightarrow \mathbb{R}_+$ satisfying

$$0 < \mathbb{E}[u(\mathbf{X}); P_{\mathbf{X};\theta}] < \infty.$$

A transform $T_u : P_{\mathbf{X};\theta} \rightarrow Q_{\mathbf{X};\theta}^{(u)}$ is defined as:

$$T_u [P_{\mathbf{X};\theta}] (A) = Q_{\mathbf{X};\theta}^{(u)} (A) \triangleq \int_A \varphi_u(\mathbf{x}; \boldsymbol{\theta}) dP_{\mathbf{X};\theta}(\mathbf{x}),$$

where

$$\varphi_u(\mathbf{x}; \boldsymbol{\theta}) \triangleq \frac{u(\mathbf{x})}{\mathbb{E}[u(\mathbf{X}); P_{\mathbf{X};\theta}]}.$$

The function $u(\cdot)$ is called the *MT-function*.

Probability Measure Transform

The measure transformed mean and covariance

$$\boldsymbol{\mu}_{\mathbf{X}}^{(u)}(\boldsymbol{\theta}) = \mathbb{E}[\mathbf{X}\varphi_u(\mathbf{X}; \boldsymbol{\theta}); P_{\mathbf{X};\boldsymbol{\theta}}]$$

$$\boldsymbol{\Sigma}_{\mathbf{X}}^{(u)}(\boldsymbol{\theta}) = \mathbb{E}[\mathbf{X}\mathbf{X}^H \varphi_u(\mathbf{X}; \boldsymbol{\theta}); P_{\mathbf{X};\boldsymbol{\theta}}] - \boldsymbol{\mu}_{\mathbf{X}}^{(u)}(\boldsymbol{\theta}) \boldsymbol{\mu}_{\mathbf{X}}^{(u)H}(\boldsymbol{\theta})$$

where

$$\varphi_u(\mathbf{x}; \boldsymbol{\theta}) \triangleq \frac{u(\mathbf{x})}{\mathbb{E}[u(\mathbf{X}); P_{\mathbf{X};\boldsymbol{\theta}}]} = \frac{dQ_{\mathbf{X};\boldsymbol{\theta}}^{(u)}}{dP_{\mathbf{X};\boldsymbol{\theta}}}$$

Conclusion

- ▶ The mean and covariance under $Q_{\mathbf{X};\boldsymbol{\theta}}^{(u)}$ can be estimated using **only** samples from $P_{\mathbf{X};\boldsymbol{\theta}}$.
- ▶ $u(\mathbf{x})$ **non-constant & analytic** \Rightarrow the mean and covariance under $Q_{\mathbf{X};\boldsymbol{\theta}}^{(u)}$ **involve higher-order statistical moments** of $P_{\mathbf{X};\boldsymbol{\theta}}$.

Probability Measure Transform

Proposition (Consistent empirical MT mean and covariance)

Let \mathbf{X}_n , $n = 1, \dots, N$ denote a sequence of i.i.d. samples from $P_{\mathbf{X};\theta}$, and define the empirical mean and covariance estimates:

$$\hat{\boldsymbol{\mu}}_{\mathbf{X}}^{(u)} \triangleq \sum_{n=1}^N \mathbf{X}_n \hat{\varphi}_u(\mathbf{X}_n)$$

$$\hat{\boldsymbol{\Sigma}}_{\mathbf{X}}^{(u)} \triangleq \sum_{n=1}^N \mathbf{X}_n \mathbf{X}_n^H \hat{\varphi}_u(\mathbf{X}_n) - \hat{\boldsymbol{\mu}}_{\mathbf{X}}^{(u)} \hat{\boldsymbol{\mu}}_{\mathbf{X}}^{(u)H}$$

where $\hat{\varphi}_u(\mathbf{X}_n) \triangleq \frac{u(\mathbf{X}_n)}{\sum_{n=1}^N u(\mathbf{X}_n)}$. If

$$\mathbb{E} \left[\|\mathbf{X}\|^2 u(\mathbf{X}); P_{\mathbf{X}} \right] < \infty,$$

then $\hat{\boldsymbol{\mu}}_{\mathbf{X}}^{(u)} \xrightarrow{w.p.1} \boldsymbol{\mu}_{\mathbf{X}}^{(u)}(\boldsymbol{\theta})$ and $\hat{\boldsymbol{\Sigma}}_{\mathbf{X}}^{(u)} \xrightarrow{w.p.1} \boldsymbol{\Sigma}_{\mathbf{X}}^{(u)}(\boldsymbol{\theta})$ as $N \rightarrow \infty$.

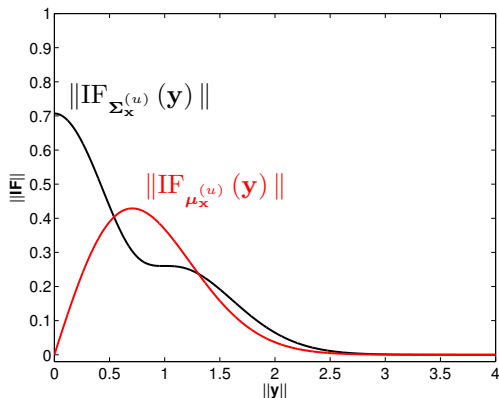
Probability Measure Transform

Proposition (Robustness to outliers)

If the MT-function $u(\mathbf{x})$ and $u(\mathbf{x})\|\mathbf{x}\|^2$ are bounded, then the *influence functions* [Hampel, 1974] of $\hat{\mu}_{\mathbf{x}}^{(u)}$ and $\hat{\Sigma}_{\mathbf{x}}^{(u)}$ are *bounded*.

Remark

Condition is satisfied when $u(\mathbf{x}) \in$ Gaussian family.



Measure Transformed Gaussian QML Estimator

The MT-GQMLE

- ▶ Minimize the *empirical KLD* between $Q_{\mathbf{X};\theta_0}^{(u)}$ and a Gaussian measure $\Phi_{\mathbf{X};\theta}^{(u)}$ with mean $\boldsymbol{\mu}_{\mathbf{X}}^{(u)}(\boldsymbol{\theta})$ and covariance $\boldsymbol{\Sigma}_{\mathbf{X}}^{(u)}(\boldsymbol{\theta})$.
- ▶ Amounts to maximization of

$$J_u(\boldsymbol{\theta}) \triangleq -D_{\text{LD}} \left[\hat{\boldsymbol{\Sigma}}_{\mathbf{X}}^{(u)} \parallel \boldsymbol{\Sigma}_{\mathbf{X}}^{(u)}(\boldsymbol{\theta}) \right] - \left\| \hat{\boldsymbol{\mu}}_{\mathbf{X}}^{(u)} - \boldsymbol{\mu}_{\mathbf{X}}^{(u)}(\boldsymbol{\theta}) \right\|_{\left(\boldsymbol{\Sigma}_{\mathbf{X}}^{(u)}(\boldsymbol{\theta}) \right)^{-1}}^2$$

- ▶ The MT-QMLE:

$$\hat{\boldsymbol{\theta}}_u = \arg \max_{\boldsymbol{\theta} \in \Theta} J_u(\boldsymbol{\theta})$$

Measure Transformed Gaussian QML Estimator

Theorem (Strong consistency of $\hat{\theta}_u$)

Given a sequence of N i.i.d. samples from $P_{\mathbf{X};\theta_0}$. Assume that the following conditions are satisfied:

1. The parameter space Θ is compact.
2. $\mu_{\mathbf{X}}^{(u)}(\theta_0) \neq \mu_{\mathbf{X}}^{(u)}(\theta)$ or $\Sigma_{\mathbf{X}}^{(u)}(\theta_0) \neq \Sigma_{\mathbf{X}}^{(u)}(\theta) \forall \theta \neq \theta_0$.
3. $\Sigma_{\mathbf{X}}^{(u)}(\theta)$ is non-singular $\forall \theta \in \Theta$.
4. $\mu_{\mathbf{X}}^{(u)}(\theta)$ and $\Sigma_{\mathbf{X}}^{(u)}(\theta)$ are continuous in Θ .
5. $E \left[\|\mathbf{X}\|^2 u(\mathbf{X}); P_{\mathbf{X};\theta_0} \right] < \infty$.

Then,

$$\hat{\theta}_u \xrightarrow{w.p. 1} \theta_0 \text{ as } N \rightarrow \infty$$

Measure Transformed Gaussian QML Estimator

Theorem (Asymptotic normality and unbiasedness of $\hat{\theta}_u$)

Given a sequence of N i.i.d. samples from $P_{\mathbf{X};\theta_0}$. Assume that the following conditions are satisfied:

1. $\hat{\theta}_u \xrightarrow{P} \theta_0$ as $N \rightarrow \infty$.
2. θ_0 lies in the interior of Θ which is assumed to be compact.
3. $\mu_{\mathbf{X}}^{(u)}(\theta)$, $\Sigma_{\mathbf{X}}^{(u)}(\theta)$ are twice continuously differentiable in Θ .
4. $E[u^2(\mathbf{X}); P_{\mathbf{X};\theta_0}] < \infty$ and $E[\|\mathbf{X}\|^4 u^2(\mathbf{X}); P_{\mathbf{X};\theta_0}] < \infty$.

Then,

$$\hat{\theta}_u \stackrel{a}{\sim} \mathcal{N}(\theta_0, \mathbf{C}_u(\theta_0))$$

Measure Transformed Gaussian QML Estimator

Asymptotic MSE

$$\mathbf{C}_u(\boldsymbol{\theta}_0) = N^{-1} \mathbf{F}_u^{-1}(\boldsymbol{\theta}_0) \mathbf{G}_u(\boldsymbol{\theta}_0) \mathbf{F}_u^{-1}(\boldsymbol{\theta}_0)$$

where

$$\mathbf{G}_u(\boldsymbol{\theta}) \triangleq \mathbb{E} \left[u^2(\mathbf{X}) \boldsymbol{\psi}_u(\mathbf{X}; \boldsymbol{\theta}) \boldsymbol{\psi}_u^T(\mathbf{X}; \boldsymbol{\theta}); P_{\mathbf{X}; \boldsymbol{\theta}_0} \right]$$

$$\boldsymbol{\psi}_u(\mathbf{X}; \boldsymbol{\theta}) \triangleq \nabla_{\boldsymbol{\theta}_0} \log \phi^{(u)}(\mathbf{X}; \boldsymbol{\theta})$$

$$\mathbf{F}_u(\boldsymbol{\theta}) \triangleq -\mathbb{E} \left[u(\mathbf{X}) \boldsymbol{\Gamma}_u(\mathbf{X}; \boldsymbol{\theta}); P_{\mathbf{X}; \boldsymbol{\theta}_0} \right]$$

$$\boldsymbol{\Gamma}_u(\mathbf{X}; \boldsymbol{\theta}) \triangleq \nabla_{\boldsymbol{\theta}}^2 \log \phi^{(u)}(\mathbf{X}; \boldsymbol{\theta})$$

Measure Transformed Gaussian QML Estimator

Proposition (Relation to the CRLB)

$$\mathbf{C}_u(\boldsymbol{\theta}_0) \succeq \mathbf{CRLB}(\boldsymbol{\theta}_0)$$

where equality holds if and only if

$$\nabla_{\boldsymbol{\theta}} \log f(\mathbf{X}; \boldsymbol{\theta})|_{\boldsymbol{\theta}=\boldsymbol{\theta}_0} = \mathbf{I}_{\text{FIM}}(\boldsymbol{\theta}_0) \mathbf{F}_u^{-1}(\boldsymbol{\theta}_0) \boldsymbol{\psi}_u(\mathbf{X}; \boldsymbol{\theta}_0) u(\mathbf{X}) \quad w.p. 1$$

Remark

$P_{\mathbf{X};\boldsymbol{\theta}_0}$ Gaussian \Rightarrow Condition is satisfied only for $u(\mathbf{x}) = \text{const}$

Conclusion

$P_{\mathbf{X};\boldsymbol{\theta}_0}$ Gaussian & $u(\mathbf{x}) \neq \text{const} \Rightarrow \mathbf{C}_u(\boldsymbol{\theta}_0) \succ \mathbf{CRLB}(\boldsymbol{\theta}_0)$

Measure Transformed Gaussian QML Estimator

Theorem (Empirical asymptotic MSE)

Define the empirical asymptotic MSE:

$$\hat{\mathbf{C}}_u(\hat{\boldsymbol{\theta}}_u) \triangleq N^{-1} \hat{\mathbf{F}}_u^{-1}(\hat{\boldsymbol{\theta}}_u) \hat{\mathbf{G}}_u(\hat{\boldsymbol{\theta}}_u) \hat{\mathbf{F}}_u^{-1}(\hat{\boldsymbol{\theta}}_u)$$

where

$$\hat{\mathbf{G}}_u(\boldsymbol{\theta}) \triangleq N^{-1} \sum_{n=1}^N u^2(\mathbf{X}_n) \boldsymbol{\psi}_u(\mathbf{X}_n; \boldsymbol{\theta}) \boldsymbol{\psi}_u^T(\mathbf{X}_n; \boldsymbol{\theta})$$

$$\hat{\mathbf{F}}_u(\boldsymbol{\theta}) \triangleq -N^{-1} \sum_{n=1}^N u(\mathbf{X}_n) \boldsymbol{\Gamma}_u(\mathbf{X}_n; \boldsymbol{\theta})$$

Furthermore, assume that the following conditions are satisfied:

1. $\hat{\boldsymbol{\theta}}_u \xrightarrow{P} \boldsymbol{\theta}_0$ as $N \rightarrow \infty$.
2. $\boldsymbol{\mu}_X^{(u)}(\boldsymbol{\theta})$, $\boldsymbol{\Sigma}_X^{(u)}(\boldsymbol{\theta})$ are twice continuously differentiable in $\boldsymbol{\Theta}$.
3. $E[u^2(\mathbf{X}); P_{\mathbf{X}; \boldsymbol{\theta}_0}] < \infty$ and $E[\|\mathbf{X}\|^4 u^2(\mathbf{X}); P_{\mathbf{X}; \boldsymbol{\theta}_0}] < \infty$.

Then,

$$N \|\hat{\mathbf{C}}_u(\hat{\boldsymbol{\theta}}_u) - \mathbf{C}_u(\boldsymbol{\theta}_0)\| \xrightarrow{P} 0 \text{ as } N \rightarrow \infty.$$

Measure Transformed Gaussian QML Estimator

Optimal choice of the MT-function

- ▶ Specify the MT-function within some parametric family

$$\{u(\mathbf{X}; \boldsymbol{\omega}), \boldsymbol{\omega} \in \boldsymbol{\Omega} \subseteq \mathbb{C}^r\}$$

- ▶ In order to gain *robustness against outliers* the *Gaussian family* would be a good choice
- ▶ An optimal choice of the MT-function parameter $\boldsymbol{\omega}$ would be this that minimizes the empirical asymptotic MSE

$$\boldsymbol{\omega}_{\text{opt}} \triangleq \arg \min_{\boldsymbol{\omega} \in \boldsymbol{\Omega}} \hat{\mathbf{C}}_u \left(\hat{\boldsymbol{\theta}}_u(\boldsymbol{\omega}); \boldsymbol{\omega} \right)$$

Application

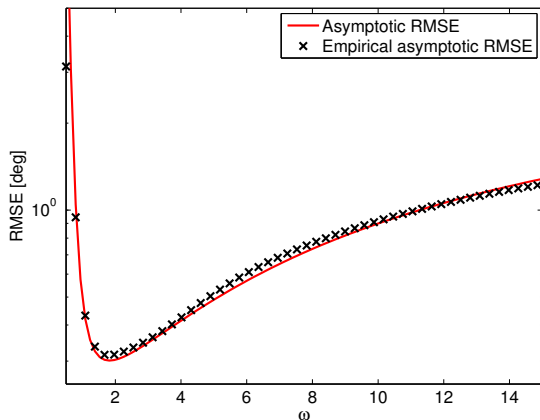
Robust direction finding in heavy-tailed clutter

$$\mathbf{X}_n = S_n \mathbf{a}(\theta_0) + \mathbf{W}_n, \quad n = 1, \dots, N$$

- ▶ S_n : emitted signal with *unknown* symmetric distribution
- ▶ \mathbf{W}_n : clutter with *unknown* spherically symmetric distribution
- ▶ S_n & \mathbf{W}_n statistically independent and first-order stationary
- ▶ **Gaussian MT-function:** $u(\mathbf{x}; \omega) \triangleq \exp(-\|\mathbf{x}\|^2/\omega^2)$
 - ▶ MT-Mean: $\boldsymbol{\mu}_{\mathbf{X}}^{(u)}(\theta; \omega) = \mathbf{0}$
 - ▶ MT-Covariance: $\boldsymbol{\Sigma}_{\mathbf{X}}^{(u)}(\theta; \omega) = r_S(\omega) \mathbf{a}(\theta) \mathbf{a}^H(\theta) + r_W(\omega) \mathbf{I}$
- ▶ **MT-GQMLE:** $\hat{\theta}_u(\omega) = \arg \max_{\theta \in \Theta} \mathbf{a}^H(\theta) \hat{\mathbf{C}}_{\mathbf{X}}^{(u)}(\omega) \mathbf{a}(\theta)$

Application

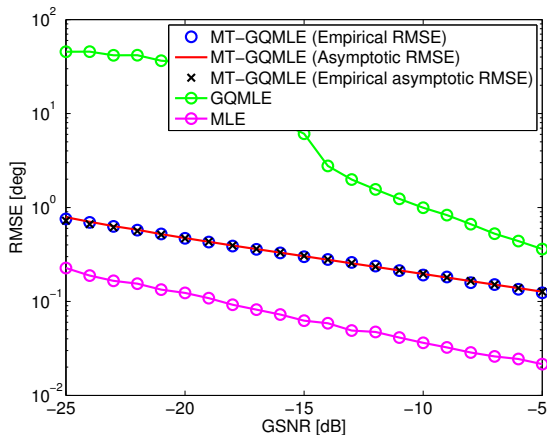
Robust direction finding in heavy-tailed clutter



- ▶ BPSK signal, 4-element ULA, $\theta = 30^\circ$
- ▶ Impulsive **K-distributed** clutter
- ▶ $N = 3000$ snapshots, GSNR = -15 [dB]

Application

Robust direction finding in heavy-tailed clutter



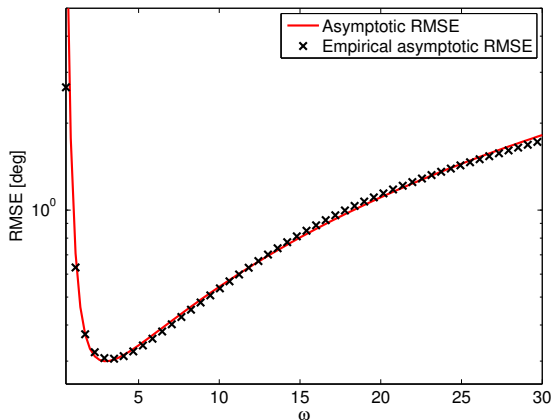
- ▶ BPSK signal, 4-element ULA, $\theta = 30^\circ$
- ▶ Impulsive **K-distributed** clutter
- ▶ $N = 3000$ snapshots

Summary

- ▶ A new estimator, called MT-GQMLE, was derived by applying a transform to the probability distribution of the data.
- ▶ By specifying the MT-function in the Gaussian family, the MT-GQMLE was applied to robust direction finding.
- ▶ Exploration of other MT-functions may result in additional estimators in this class that have different useful properties.

Application

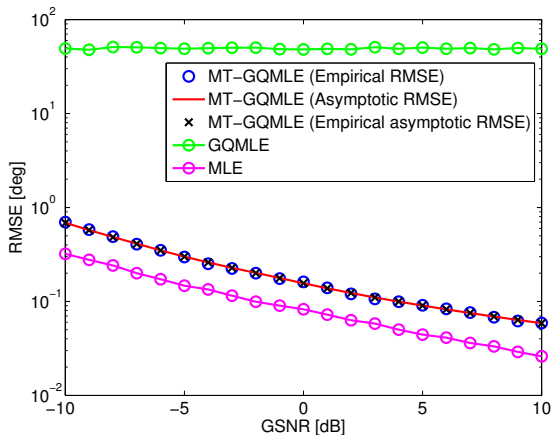
Robust direction finding in heavy-tailed clutter



- ▶ BPSK signal, 4-element ULA, $\theta = 30^\circ$
- ▶ Impulsive **Cauchy** clutter
- ▶ $N = 3000$ snapshots, $\text{GSNR} = -5$ [dB]

Application

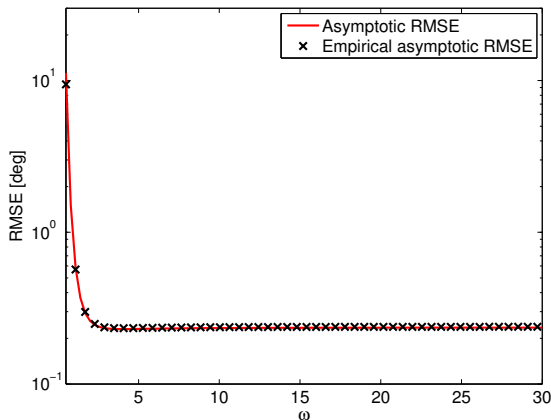
Robust direction finding in heavy-tailed clutter



- ▶ BPSK signal, 4-element ULA, $\theta = 30^\circ$
- ▶ Impulsive **Cauchy** clutter
- ▶ $N = 3000$ snapshots

Application

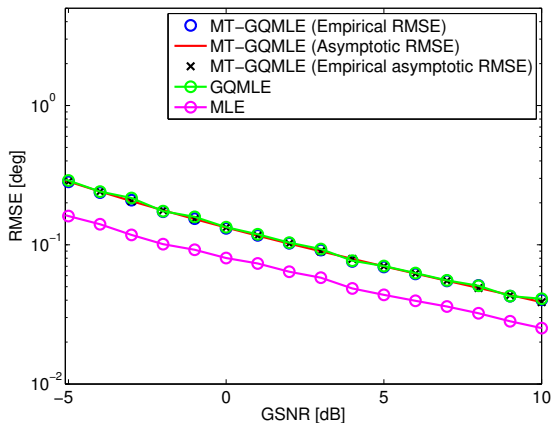
Robust direction finding in Gaussian clutter



- ▶ BPSK signal, 4-element ULA, $\theta = 30^\circ$
- ▶ Gaussian clutter
- ▶ $N = 3000$ snapshots, GSNR = -5 [dB]

Application

Robust direction finding in Gaussian clutter



- ▶ BPSK signal, 4-element ULA, $\theta = 30^\circ$
- ▶ Gaussian clutter
- ▶ $N = 3000$ snapshots