

Plug-in Measure-Transformed Quasi Likelihood Ratio Test for Random Signal Detection

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Outline

- ▶ Detection problem formulation
- ▶ Existing detectors
- ▶ Proposed detector
- ▶ Simulation studies
- ▶ Summary

Detection Problem Formulation

$$H_0 : \begin{cases} \mathbf{X}_n = \mathbf{W}_n, & n = 1, \dots, N \\ \mathbf{Y}_m = \mathbf{W}_m^{(s)}, & m = 1, \dots, M \end{cases}$$
$$H_1 : \begin{cases} \mathbf{X}_n = S_n \mathbf{a} + \mathbf{W}_n, & n = 1, \dots, N \\ \mathbf{Y}_m = \mathbf{W}_m^{(s)}, & m = 1, \dots, M, \end{cases}$$

- ▶ $\{\mathbf{X}_n \in \mathbb{C}^p\}, \{\mathbf{Y}_m \in \mathbb{C}^p\}$: Primary and secondary data
- ▶ $\{S_n \in \mathbb{C}\}$: i.i.d. zero-mean signal with *unknown* distribution
- ▶ $\mathbf{a} \in \mathbb{C}^p$: *Known* steering vector
- ▶ $\{\mathbf{W}_n\}, \{\mathbf{W}_m^{(s)}\}$: i.i.d. zero-mean symmetrically distributed *homogeneous* noise processes with *unknown* distribution
- ▶ $\{S_n\}, \{\mathbf{W}_n\}$ and $\{\mathbf{W}_m^{(s)}\}$ are mutually independent

Existing detectors

Gauss-Gauss detector [*Jin & Friedlander 2005*]

- ▶ GLRT-based detector
- ▶ Assumes normally distributed signal and noise
- ▶ Simple implementation and tractable performance analysis
- ▶ Sensitive to deviation from normality (e.g., in the case of heavy-tailed noise that produces outliers)

Existing detectors

CG-GLRT [*Gerlach 1999, Shuai et. al. 2010*]

- ▶ Assumes elliptical compound-Gaussian noise
- ▶ Resilient against heavy-tailed noise outliers
- ▶ Iterative ML-estimation of the noise scatter matrix
 - ▶ Converges under some regularity conditions
 - ▶ Each iteration involves matrix inversion
 - ▶ Does not reject large norm outliers
- ▶ Can be sensitive to non-elliptical noise

Proposed Detector

General operation principle

Selects a Gaussian model that best empirically fits a **transformed** probability distribution of the data

Advantages

- ▶ Resilient to outliers (rejects large-norm outliers)
- ▶ Involves higher-order statistical moments
- ▶ Significant mitigation of the model mismatch effect
- ▶ Computational and implementation simplicity

Probability Measure Transform

Definition

Given a non-negative function $u : \mathcal{X} \rightarrow \mathbb{R}_+$ satisfying

$$0 < \mathbb{E}[u(\mathbf{X}); P_{\mathbf{X};H}] < \infty.$$

A transform $T_u : P_{\mathbf{X};H} \rightarrow Q_{\mathbf{X};H}^{(u)}$ is defined as:

$$T_u[P_{\mathbf{X};H}](A) = Q_{\mathbf{X};H}^{(u)}(A) \triangleq \int_A \varphi_u(\mathbf{x}; H) dP_{\mathbf{X};H}(\mathbf{x}),$$

where

$$\varphi_u(\mathbf{x}; H) \triangleq \frac{u(\mathbf{x})}{\mathbb{E}[u(\mathbf{X}); P_{\mathbf{X};H}]}.$$

The function $u(\cdot)$ is called the *MT-function*.

Probability Measure Transform

The measure transformed mean and covariance

$$\boldsymbol{\mu}_{\mathbf{X};H}^{(u)} = \mathbb{E} [\mathbf{X} \varphi_u (\mathbf{X}; H); P_{\mathbf{X};H}]$$

$$\boldsymbol{\Sigma}_{\mathbf{X};H}^{(u)} = \mathbb{E} [\mathbf{X} \mathbf{X}^H \varphi_u (\mathbf{X}; H); P_{\mathbf{X};H}] - \boldsymbol{\mu}_{\mathbf{X};H}^{(u)} \boldsymbol{\mu}_{\mathbf{X};H}^{(u)H}$$

where

$$\varphi_u (\mathbf{x}; H) \triangleq \frac{u (\mathbf{x})}{\mathbb{E} [u (\mathbf{X}); P_{\mathbf{X};H}]} = \frac{dQ_{\mathbf{X};H}^{(u)}}{dP_{\mathbf{X};H}}$$

Conclusion

- ▶ The mean and covariance under $Q_{\mathbf{X};H}^{(u)}$ can be estimated using **only** samples from $P_{\mathbf{X};H}$.
- ▶ $u (\mathbf{x})$ **non-constant & analytic** \Rightarrow the mean and covariance under $Q_{\mathbf{X};H}^{(u)}$ **involve higher-order statistical moments** of $P_{\mathbf{X};H}$.

Probability Measure Transform

Proposition (Consistent empirical MT mean and covariance)

Let \mathbf{X}_n , $n = 1, \dots, N$ denote a sequence of i.i.d. samples from $P_{\mathbf{X};H}$, and define the empirical mean and covariance estimates:

$$\hat{\boldsymbol{\mu}}_{\mathbf{X}}^{(u)} \triangleq \sum_{n=1}^N \mathbf{X}_n \hat{\varphi}_u(\mathbf{X}_n)$$

$$\hat{\boldsymbol{\Sigma}}_{\mathbf{X}}^{(u)} \triangleq \sum_{n=1}^N \mathbf{X}_n \mathbf{X}_n^H \hat{\varphi}_u(\mathbf{X}_n) - \hat{\boldsymbol{\mu}}_{\mathbf{X}}^{(u)} \hat{\boldsymbol{\mu}}_{\mathbf{X}}^{(u)H}$$

where $\hat{\varphi}_u(\mathbf{X}_n) \triangleq \frac{u(\mathbf{X}_n)}{\sum_{n=1}^N u(\mathbf{X}_n)}$. If

$$\mathbb{E} \left[\|\mathbf{X}\|^2 u(\mathbf{X}); P_{\mathbf{X};H} \right] < \infty,$$

then $\hat{\boldsymbol{\mu}}_{\mathbf{X}}^{(u)} \xrightarrow{w.p.1} \boldsymbol{\mu}_{\mathbf{X};H}^{(u)}$ and $\hat{\boldsymbol{\Sigma}}_{\mathbf{X}}^{(u)} \xrightarrow{w.p.1} \boldsymbol{\Sigma}_{\mathbf{X};H}^{(u)}$ as $N \rightarrow \infty$.

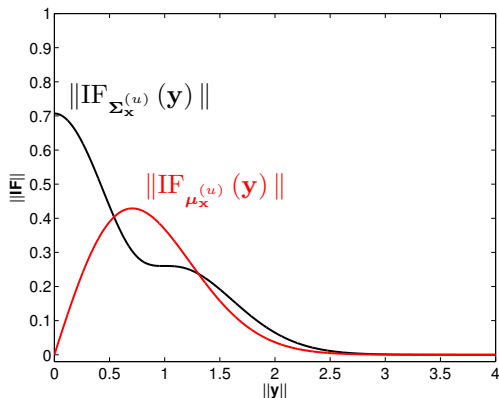
Probability Measure Transform

Proposition (Robustness to outliers)

If the MT-function $u(\mathbf{x})$ and $u(\mathbf{x})\|\mathbf{x}\|^2$ are bounded, then the *influence functions* [Hampel, 1974] of $\hat{\mu}_{\mathbf{x}}^{(u)}$ and $\hat{\Sigma}_{\mathbf{x}}^{(u)}$ are *bounded*.

Remark

Condition is satisfied when $u(\mathbf{x}) \in$ Gaussian family.



Measure Transformed Gaussian Quasi LRT

MT-GQLRT [*Todros & Hero, 2016*]

Compares the *empirical KLDs* between $Q_{\mathbf{X};H}^{(u)}$ and two Gaussian measures $\Phi(\boldsymbol{\mu}_{\mathbf{X};H_0}^{(u)}, \boldsymbol{\Sigma}_{\mathbf{X};H_0}^{(u)})$ and $\Phi(\boldsymbol{\mu}_{\mathbf{X};H_1}^{(u)}, \boldsymbol{\Sigma}_{\mathbf{X};H_1}^{(u)})$

$$\begin{aligned} T_u &= D_{\text{LD}} \left[\hat{\boldsymbol{\Sigma}}_{\mathbf{X}}^{(u)} \parallel \boldsymbol{\Sigma}_{\mathbf{X};H_0}^{(u)} \right] + \left\| \hat{\boldsymbol{\mu}}_{\mathbf{X}}^{(u)} - \boldsymbol{\mu}_{\mathbf{X};H_0}^{(u)} \right\|_{\left(\boldsymbol{\Sigma}_{\mathbf{X};H_0}^{(u)} \right)^{-1}}^2 \\ &- D_{\text{LD}} \left[\hat{\boldsymbol{\Sigma}}_{\mathbf{X}}^{(u)} \parallel \boldsymbol{\Sigma}_{\mathbf{X};H_1}^{(u)} \right] - \left\| \hat{\boldsymbol{\mu}}_{\mathbf{X}}^{(u)} - \boldsymbol{\mu}_{\mathbf{X};H_1}^{(u)} \right\|_{\left(\boldsymbol{\Sigma}_{\mathbf{X};H_1}^{(u)} \right)^{-1}}^2 \stackrel{H_1}{\underset{H_0}{\gtrless}} \tau, \end{aligned}$$

Measure Transformed Gaussian Quasi LRT

MT-GQLRT for the considered detection problem

- ▶ Class of MT-functions:

$$\left\{ u(\mathbf{x}) = v\left(\mathbf{P}_{\mathbf{a}}^{\perp} \mathbf{x}\right), v: \mathbb{C}^p \rightarrow \mathbb{R}_+, v(\mathbf{x}) = v(-\mathbf{x}) \right\}$$

- ▶ MT-mean and MT-covariance under H_0 and H_1 :

$$\boldsymbol{\mu}_{\mathbf{x}; H_k}^{(u)} = \mathbf{0}, \quad k = 0, 1$$

$$\boldsymbol{\Sigma}_{\mathbf{x}; H_0}^{(u)} = \boldsymbol{\Sigma}_{\mathbf{w}}^{(u)} \quad \text{and} \quad \boldsymbol{\Sigma}_{\mathbf{x}; H_1}^{(u)} = \sigma_S^2 \mathbf{a} \mathbf{a}^H + \boldsymbol{\Sigma}_{\mathbf{w}}^{(u)}$$

- ▶ Equivalent test statistic:

$$T'_u = \mathbf{a}^H \left(\boldsymbol{\Sigma}_{\mathbf{w}}^{(u)} \right)^{-1} \hat{\mathbf{C}}_{\mathbf{x}}^{(u)} \left(\boldsymbol{\Sigma}_{\mathbf{w}}^{(u)} \right)^{-1} \mathbf{a}$$

Plug-in Measure Transformed Gaussian Quasi LRT

Plug-in MT-GQLRT

Plug-in the empirical MT-covariance of the noise obtained from noise-only secondary data $\{\mathbf{Y}_m\}$

$$T_u'' \triangleq \mathbf{a}^H \left(\hat{\Sigma}_{\mathbf{Y}}^{(u)} \right)^{-1} \hat{\mathbf{C}}_{\mathbf{X}}^{(u)} \left(\hat{\Sigma}_{\mathbf{Y}}^{(u)} \right)^{-1} \mathbf{a} \underset{H_0}{\overset{H_1}{>}} t,$$

Theorem (Asymptotic normality)

Assume that

1. $\sigma_S^2 > 0$, $\Sigma_{\mathbf{W}}^{(u)}$ is non-singular
2. $E[u^2(\mathbf{X}); P_{\mathbf{X};H}]$, $E[\|\mathbf{X}\|^4 u^2(\mathbf{X}); P_{\mathbf{X};H}]$ are finite

Then

$$\sqrt{N} \left(T_u'' - \eta_H^{(u)} \right) \xrightarrow[N \rightarrow \infty]{D} \mathcal{N} \left(0, \lambda_H^{(u)} \right) \text{ for } H = H_0, H_1$$

Plug-in Measure Transformed Gaussian Quasi LRT

Threshold determination

$$t = \hat{\eta}_{H_0}^{(u)} + Q^{-1}(\alpha) \sqrt{\hat{\lambda}_{H_0}^{(u)}},$$

- ▶ Guarantees asymptotic CFAR = α

- ▶ $\hat{\eta}_{H_0}^{(u)} \triangleq \mathbf{a}^H \left(\hat{\Sigma}_{\mathbf{Y}}^{(u)} \right)^{-1} \mathbf{a}$

- ▶ $\hat{\lambda}_{H_0}^{(u)} \triangleq \frac{M}{N} \sum_{m=1}^M \hat{\varphi}_u^2(\mathbf{Y}_m) \left(\left| \mathbf{a}^H \left(\hat{\Sigma}_{\mathbf{Y}}^{(u)} \right)^{-1} \mathbf{Y}_m \right|^2 - \hat{\eta}_{H_0}^{(u)} \right)^2$

Plug-in Measure Transformed Gaussian Quasi LRT

Selection of the MT-function to induce outlier resilience

- ▶ Gaussian MT-function:

$$u_G(\mathbf{x}; \omega) = \exp\left(-\|\mathbf{P}_a^\perp \mathbf{x}\|^2 / \omega^2\right)$$

- ▶ Define the **Asymptotic Relative Local Power Sensitivity** to change in σ_S^2 (under nominal Gaussian distribution):

$$R(\omega, \Sigma_{\mathbf{w}}) \triangleq \frac{\partial \beta_{u_G}}{\partial \sigma_S^2} / \frac{\partial \beta_{LRT}}{\partial \sigma_S^2} \Big|_{\sigma_S^2=0} = \frac{\sqrt{N}G(\omega, \Sigma_{\mathbf{w}}) + Q^{-1}(\alpha)}{\sqrt{N} + Q^{-1}(\alpha)}$$

- ▶ Select the lowest ω such that $R(\omega, \Sigma_{\mathbf{w}}) \geq r$, $0 \ll r < 1$
- ▶ In practice $\Sigma_{\mathbf{w}}$ is replaced by:

$$\hat{\Sigma}_{\mathbf{w}}(\omega) = \left(\hat{\Sigma}_{\mathbf{Y}}^{(u_G)}(\omega) - \mathbf{P}_a^\perp / \omega^2\right)^{-1}$$

Simulation Studies

Setup

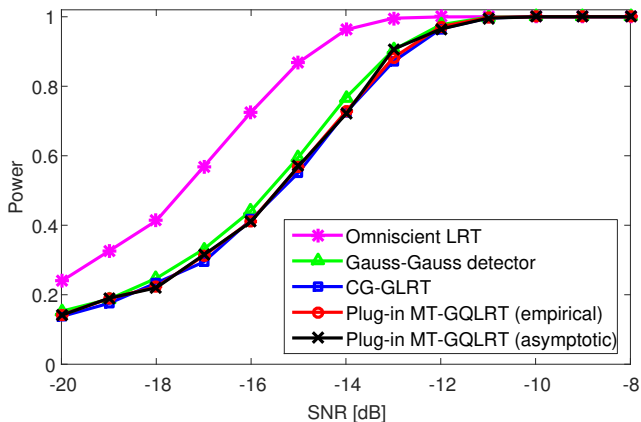
- ▶ BPSK signal with variance σ_S^2
- ▶ Steering vector: $\mathbf{a} \triangleq \frac{1}{\sqrt{p}} [1, e^{-i\pi \sin(\theta)}, \dots, e^{-i\pi(p-1) \sin(\theta)}]^T$,
 $p = 8, \theta = 60^\circ$
- ▶ Sample size in primary and secondary data: $N = 1000$,
 $M = 5000$
- ▶ False alarm rate $\alpha = 0.05$
- ▶ Asymptotic Relative Local Power Sensitivity: $r = 0.9$

Simulation Studies

Detection in Gaussian noise

Zero-mean Gaussian noise with Toeplitz structured covariance Σ :

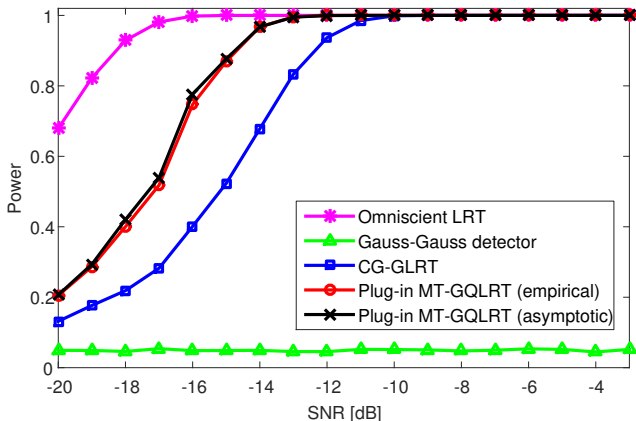
$$[\Sigma]_{i,j} \triangleq \begin{cases} \sigma_{\mathbf{w}}^2 b^{j-i}, & i \leq j \\ \sigma_{\mathbf{w}}^2 (b^{i-j})^*, & i > j \end{cases}, \quad |b| < 1,$$



Simulation Studies

Detection in elliptical CG noise

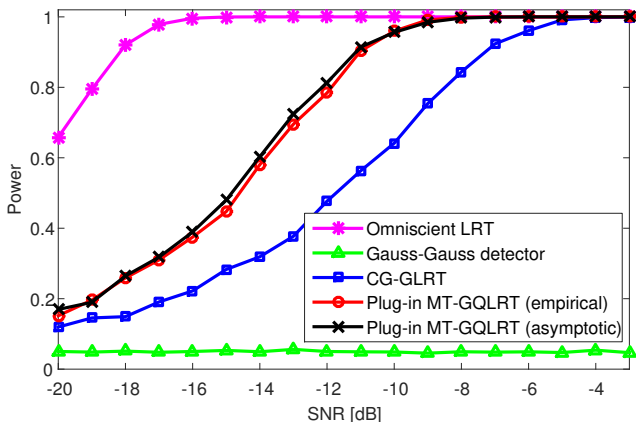
Elliptical t -distributed noise with $\lambda = 0.2$ degrees of freedom and dispersion matrix Σ



Simulation Studies

Detection in non-elliptical noise

Elliptical t -distributed noise + BPSK interference (SIR = -5 [dB])



Conclusions

- ▶ Plug-in MT-GQLRT for detection of a random signal that lies on a known rank-one subspace
- ▶ Simple implementation and tractable performance analysis
- ▶ Less sensitive to model mismatch as compared to other model based detectors
- ▶ Extension to non-homogeneous noise environment